

Liquidity traps: how to avoid them and how to escape them

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Abstract

An economy is in a liquidity trap when monetary policy cannot influence either real or nominal variables of interest. A necessary condition for this is that the short nominal interest rate is constrained by its lower bound, typically zero. The paper develops a small analytical model to show how an economy can get into a liquidity trap, how it can avoid getting into one and how it can get out. We also consider the empirical likelihood of the UK economy hitting the zero nominal rate bound, by investigating the relationship between the level of the short nominal interest rate and its volatility. The empirical evidence on this issue is mixed. To reduce the risk of falling into a liquidity trap, the authorities have two options. The first is to raise the inflation target. The second is to lower the zero nominal interest rate floor. This second option involves paying negative interest on government 'bearer bonds' – coin and currency – ie 'taxing money', as advocated by Gesell. Once in a liquidity trap, there are two means of escape. The first is to use expansionary fiscal policy. The second is, again, to lower the zero nominal interest rate floor. There are likely to be significant shoe leather costs associated with any scheme to tax currency.

1. Introduction

The liquidity trap used to be a standard topic in macro textbooks, but disappeared in the 1970s. Because of recent developments in Japan, liquidity traps are a hot topic again. The economy is said to be in a liquidity trap when the ability to use monetary policy to stimulate demand has vanished. The conditions that have to be satisfied for monetary policy to fail to affect both real and nominal variables depend on one's view on the monetary transmission mechanism. A necessary condition for monetary policy ineffectiveness is that monetary policy cannot affect the joint distribution of real and nominal rates of return on financial and real assets.⁽¹⁾ In very simple closed economy IS-LM-type models, where there is but one rate of return – a nominal interest rate of unspecified maturity – monetary policy is powerless when an increase in the nominal money stock cannot reduce this short nominal rate. In models with a more extended menu of financial and real assets, for monetary policy to be powerless, the yields on all non-monetary assets (short and long maturity, private and public, financial and real) must be at their lower bounds, not just the short nominal interest rate. With portfolio holders indifferent as regards the composition of their financial wealth between money and all non-money assets, changes in the supply of money cannot affect the spreads between money and non-money assets. When, as is institutionally more relevant, the short nominal interest rate is taken to be the monetary instrument, rather than some monetary aggregate, the argument is not changed in any essential way. When monetary policy also works through channels other than rates of return (say, through the availability as well as the cost of credit, or through the exchange rate), a

(1) In an open economy, the relevant rates of return would include the expected rate of depreciation of the exchange rate. If, say, domestic nominal interest rates are linked to world interest rates through an uncovered interest parity condition, monetary policy will still be ineffective whenever the demand for money becomes infinitely interest-sensitive. When international interest rate differentials can be influenced by changes in the relative supplies of non-monetary government debt instruments denominated in different currencies, there is a further monetary transmission channel.

liquidity trap is only operative if these additional channels of monetary transmission too are blocked.⁽²⁾

The textbook treatment of liquidity traps, based on Hicks's (1936) interpretation of Keynes (1936), involves the assumption that the opportunity cost of holding money is a long nominal interest rate, and that the demand for money becomes infinitely sensitive to the current value of this long nominal yield because of regressive (what we now call 'mean reverting') expectations about the future behaviour of the long nominal yield (see eg Tobin (1958) and Laidler (1993)). In most modern theories, the short (riskless) nominal interest rate on government debt is the opportunity cost of holding currency. The nominal yield on short government debt is then related to yields on other assets through equilibrium asset-pricing relationships, such as the expectations theory of the term structure of interest rates, the CAPM model or other portfolio balance models.

The modern argument assumes explicitly (and the traditional theories assumed implicitly) that the pecuniary own rate of return on money was zero, an appropriate assumption for coin and currency, although not for the liabilities of private deposit-taking institutions that make up most of the broader monetary aggregates, which now typically have positive nominal returns. With the own rate of return on currency administratively fixed at zero, a floor for the spread between the non-monetary and monetary claims becomes a floor for the nominal yields on non-monetary financial instruments.

(2) A 'helicopter drop' of money will, unlike money injected through open market purchases, have a wealth effect on private consumption, for a given distribution of rates of return. Since the essence of this part of the monetary transmission mechanism is a transfer of wealth between the public and private sectors, we consider it to be fiscal rather than monetary policy. In the rest of this paper, monetary policy is interpreted as pertaining only to the composition of the government's financial liabilities between monetary and non-monetary claims. The magnitude of the government's aggregate stock of financial liabilities is the province of intertemporal fiscal policy.

In Section 2 of the paper, we focus on the empirical issue of the likelihood that the instantaneous risk-free nominal yield, the short nominal interest rate, would reach its zero floor.⁽³⁾ We fully recognise that a zero short nominal rate is only a necessary, and not a sufficient, condition for the liquidity trap to be operative. Even in the simple analytical model of a liquidity trap developed in Section 4, monetary policy is powerless only if it cannot affect nominal yields at any maturity. While the demand for narrow money or base money in that model depends on just one opportunity cost variable, the current short nominal interest rate, aggregate demand is affected by current and anticipated future short rates. The economy is in a liquidity trap only if the entire yield curve is flat at a zero level (see also Orphanides and Wieland (1998)). In an open-economy extension of this model, the same conclusion would apply if domestic and foreign nominal interest rates were linked through an uncovered interest parity (UIP) condition. If domestic and foreign currency denominated non-monetary securities were imperfect substitutes, monetary policy might work through the exchange rate channel, even with the entire domestic yield curve flat at zero.

This liquidity trap used to be treated, in the mainstream accounts of the monetary transmission mechanism, as a theoretical curiosum without practical relevance.⁽⁴⁾ The revival of interest in the liquidity trap is not surprising.

(3) With the nominal rate on currency fixed at zero, the nominal interest rate on other financial claims can be negative if the cost of holding and storing currency exceeds that of holding and storing these other claims, because of differential tax treatment of currency and other stores of value or because of different collateralisability properties (see Porter (1999)). From late 1938 to early 1941, weekly data on auctions of new Treasury bills in the United States showed occasional negative yields. During this period, negative yields were also reported on US Treasury bonds with up to two years maturity (see Clouse *et al* (1999)). In Japan, short-term interest rates on government debt and some interbank lending became slightly negative in late 1998. These exceptions to the zero nominal interest rate floor have been rare and unimportant quantitatively. For expositional simplicity, we proceed in what follows as if the floor were equal to the pecuniary rate of return on currency, ie zero under historical practice.

(4) See eg Romer (1996), which covers the topic as half of an exercise at the end of Chapter 5, 'Traditional Keynesian Theories of Fluctuations'.

First, Japan is in a protracted economic slump. Short nominal interest rates there are near zero. Zero is the absolute nominal interest rate floor in Japan because yen notes and coin bear a zero nominal interest rate. Of course, the yields on longer-maturity government debt instruments remain positive (albeit at historically low levels), and the nominal yields on a variety of private and public financial and real assets also remain positive. While the strict conditions for a liquidity trap to be operative are therefore not satisfied, monetary policy in Japan currently appears to have a rather limited effect on aggregate demand. A number of observers have concluded that there is a liquidity trap at work (see eg Krugman (1998a,b,c,d; 1999), Ito (1998) and McKinnon and Ohno (1999)); for a view that liquidity traps are unlikely to pose a problem, see Meltzer (1999)).

Second, inflation in Euroland is around 1% per annum. The ECB's repo rate now⁽⁵⁾ is 2.5%. This raises the question as to whether a margin of 250 basis points provides enough insurance against a slump in aggregate demand. Demand could weaken to such an extent that a cut in the short nominal rate of more than 250 points would be required to boost aggregate demand sufficiently.

The monetary instrument is, almost invariably, a short nominal interest rate.⁽⁶⁾ Monetary policy impacts aggregate demand primarily through its effect on real interest rates, short and long. The transmission of monetary policy

(5) June 1999.

(6) The argument could be recast in terms of the monetary authority using some monetary aggregate as the instrument, with the short nominal interest rate on risk-free non-monetary financial claims treated as endogenous. Taking the short nominal rate as the instrument has two advantages. First, the exposition is simpler. Second, it is what central banks actually do. Changes in reserve requirements, open market operations etc, are best viewed as ways of changing the interest rate. In an open economy, the other institutionally relevant instrument of monetary policy is the nominal exchange rate. When capital mobility is limited, the short nominal interest rate and the nominal exchange rate both can be instruments of policy, at any rate in the short run.

through other real asset prices, including the real exchange rate, depends on the ability of the monetary authorities to influence real interest rates. For the monetary authority to affect real demand, changes in nominal interest rates have to be translated, at least temporarily, into changes in real interest rates. In a moderate or low-inflation environment, inflation and inflation expectations tend to move only gradually and sluggishly. This Keynesian feature of the economy gives monetary policy a temporary handle on the real economy.

If short nominal interest rates cannot fall any further, short real rates can only be pushed down through a rise in the expected rate of inflation. If the price stability gospel has been widely internalised by market participants, expected inflation is unlikely to rise to produce the required cut in real rates.

Once an economy is in such a situation, it is not possible to get out of it using the conventional monetary policy instruments, ie changes in the short nominal interest rates. Inflation expectations are not a policy instrument. Why would inflation expectations rise when monetary policy cannot stimulate demand? Conventional monetary policy advice then can only be preventive, not curative: do not get into this situation. Make sure inflation expectations (and actual inflation) are targeted at a level high enough to ensure that nominal interest rates will not hit the floor, even during periods during which aggressively expansionary monetary policy may be in order.

Of course, in a liquidity trap, expansionary fiscal policy, or any other exogenous shock to aggregate demand, is supposed to be at its most effective. There are, however, conditions under which fiscal policy cannot be used to stimulate aggregate demand. Debt-financed lump-sum tax cuts could fail to stimulate aggregate demand if there is Ricardian equivalence or debt neutrality. Alternatively, the government's creditworthiness may be so impaired that it cannot borrow. Finally, there could be external constraints – like the Maastricht Treaty or Stability and Growth Pact – on a government's ability to use deficit financing.

If Ricardian equivalence holds, a temporary increase in exhaustive public spending will, even with a balanced budget, and in virtually any model of the economy, boost aggregate demand. For this fiscal policy channel to be ineffective also, exhaustive public spending must be a direct perfect substitute for exhaustive private spending, say because public consumption is a perfect substitute for private consumption in private utility functions, and public investment is a perfect substitute for private investment in private production functions.⁽⁷⁾

2. A simple model of the liquidity trap

We model a simple, closed endowment economy with a single perishable commodity that can be consumed privately or publicly.

Households

A representative infinitely-lived, competitive consumer maximises for all $t \geq 0$ the utility functional given in (1) subject to his instantaneous flow budget identity (2), solvency constraint (3) and his initial financial wealth. We use the simplest money-in-the-direct-utility-function approach to motivate a demand for money despite it being dominated as a store of value. Instantaneous felicity therefore depends on consumption and real money balances. We define the following notation; c is real private consumption, y is real output, t is real (lump-sum taxes), M is the nominal stock of base money (currency), B is the nominal stock of short (strictly zero maturity) non-monetary debt, i is the instantaneous risk-free nominal interest rate on non-monetary debt, i_M is the instantaneous risk-free nominal interest rate on money, p is the price level in terms of money, a is the real stock of private

(7) See Buiter (1977).

financial wealth, m is the stock of real currency and b the stock of real non-monetary debt. So we have:

$$\int_t^{\infty} e^{-\mathbf{d}(v-t)} \left[\frac{1}{1+\mathbf{h}} \ln c(v) + \frac{\mathbf{h}}{1+\mathbf{h}} \ln m(v) \right] dv \quad (1)$$

$$\mathbf{h} > 0$$

$$\mathbf{d} > 0$$

$$\dot{M} + \dot{B} \equiv p(y - \mathbf{t} - c) + iB + i_M M \quad (2)$$

$$c \geq 0; M \geq 0$$

$$\lim_{v \rightarrow \infty} e^{-\int_t^v i(u) du} [M(v) + B(v)] \geq 0 \quad (3)$$

$$M(0) + B(0) = \bar{A}(0) \quad (4)$$

By definition,

$$a \equiv \frac{M + B}{P} \quad (5)$$

The household budget identity (2) can be rewritten as follows:

$$\dot{a} \equiv ra + y - \mathbf{t} - c + (i_M - i)m \quad (6)$$

where r , the instantaneous real rate of interest on non-monetary assets, is defined by:

$$r \equiv i - \mathbf{p} \quad (7)$$

and $\mathbf{p} \equiv \frac{\dot{p}}{p}$ is the instantaneous rate of inflation.

The household solvency constraint can now be rewritten as:

$$\lim_{v \rightarrow \infty} e^{-\int_t^v r(u) du} a(v) \geq 0 \quad (8)$$

and the intertemporal budget constraint for the household sector can be rewritten as:

$$\int_t^{\infty} e^{-\int_t^v r(u) du} \left[c(v) + \mathbf{t}(v) + [i(v) - i_M(v)]m(v) - y(v) \right] dv \leq a(t) \quad (9)$$

The first-order conditions for an optimum imply that the solvency constraint will hold with equality. Also,

$$\dot{c} = (r - d)c \quad (10)$$

and for $i > i_M$,

$$m = \left(\frac{h}{i - i_M} \right) c \quad (11)$$

If $i < i_M$, currency would dominate non-monetary financial assets ('bonds') as a store of value. Households would wish to take infinite long positions in money, financed by infinite short positions in non-monetary securities. The rate of return on the portfolio would be infinite. This cannot be an equilibrium.

If $i = i_d$, currency and bonds are perfect substitutes as stores of value.

With flexible prices, this will be the first-best equilibrium, characterised by satiation in real money balances. With the logarithmic utility function, satiation occurs only when the stock of money is infinite (relative to the finite consumption level). Provided the authorities provide government money and absorb private bonds in the right (infinite) amounts, this can be an equilibrium.

There is a continuum of identical consumers whose aggregate measure is normalised to 1. The individual relationships derived in this section therefore also characterise the aggregate behaviour of the consumers.

Government

The budget identity of the consolidated general government and central bank is given in (12). The level of real public consumption is denoted $g \geq 0$.

$$\dot{M} + \dot{B} \equiv iB + i_M M + p(g - t) \quad (12)$$

Again, the initial nominal value of the government's financial liabilities is predetermined, $M(0) + B(0) = \bar{A}(0)$.

This budget identity can be rewritten as:

$$\dot{a} \equiv ra + g - t + (i_M - i)m \quad (13)$$

The government solvency constraint is:

$$\lim_{v \rightarrow \infty} e^{-\int_t^v r(u) du} a(v) \leq 0 \quad (14)$$

Equations (13) and (14) imply the intertemporal government budget constraint:

$$\int_t^\infty e^{-\int_t^v r(u) du} \left[t(v) + [i(v) - i_M(v)]m(v) - g(v) \right] dv \geq a(t) \quad (15)$$

Government consumption spending is exogenous. To ensure that public consumption spending does not exceed total available resources, $\bar{y} > 0$, we therefore have to impose $g < \bar{y}$.

With a representative consumer, this model will exhibit debt neutrality or Ricardian equivalence. Without loss of generality, we therefore assume that lump-sum taxes are continuously adjusted to keep the nominal stock of public debt (monetary and non-monetary) constant, $\dot{A}(t) = 0$, $t \geq 0$, ie:

$$\begin{aligned} t &= g + ia + (i_M - i)m \\ &= g + i \frac{\bar{A}(0)}{p} + (i_M - i)m \end{aligned} \quad (16)$$

Monetary policy

The monetary authorities are again assumed to peg the nominal interest rate on currency exogenously, ie $i_M = \bar{i}_M$.

We assume in what follows that the other monetary instrument is the short nominal interest rate on bonds, rather than the level or the growth rate of the nominal money stock. There are two reasons for this. First, it simplifies the exposition. Second, it is how monetary policy is actually conducted.

The monetary authorities are assumed to follow a simplified Taylor rule for the short nominal interest rate on non-monetary financial claims, as long as this does not put the short nominal bond rate below the interest rate on currency. A standard Taylor rule for the short nominal bond rate which restricts the short nominal bond rate not to be below the short nominal rate on currency, would be:

$$i = \bar{i} + \mathbf{g}p + \mathbf{e}y \quad \text{if } \bar{i} + \mathbf{g}p + \mathbf{e}y \geq i_M$$

$$= i_M \quad \text{if } \bar{i} + \mathbf{g}p + \mathbf{e}y < i_M$$

For our purposes, all that matters is the responsiveness of the short bond rate to the inflation rate. We therefore omit feedback from the level of real GDP (or from the output gap) in what follows. The short nominal interest rate rule therefore simplifies to:

$$i = \bar{i} + \mathbf{g}p \quad \text{if } \bar{i} + \mathbf{g}p \geq i_M$$

$$= i_M \quad \text{if } \bar{i} + \mathbf{g}p < i_M \tag{17}$$

The Taylor rule is sometimes justified as a simple, ad hoc rule consistent with inflation targeting. If the target rate of inflation is constant at p^* (and equal to the steady-state rate of inflation), the intercept in the Taylor rule, \bar{i} , can be given the following interpretation:

$$\bar{i} = \mathbf{d} + (1 - \mathbf{g})p^* \tag{18}$$

This implies:

$$i = \mathbf{d} + p^* + \mathbf{g}(p - p^*) \tag{19}$$

or

$$r = \mathbf{d} + (\mathbf{g} - 1)(p - p^*) \tag{20}$$

For reasons of space, only a ‘Keynesian’ variant of the model, characterised by nominal price rigidities, is considered here.⁽⁸⁾ In this Keynesian variant, output is demand-determined, the price level and the rate of inflation are assumed to be predetermined, and the rate of inflation adjusts to the gap between actual and capacity output through the simplest kind of accelerationist Phillips curve.

$$c + g = y \quad (21)$$

$$\dot{p} = b(y - \bar{y}) \quad b > 0 \quad (22)$$

For simplicity, we assume capacity output to be exogenous and constant. The behaviour of the economy can be summarised in two first-order differential equations in the non-predetermined state variable c and the predetermined state variable p . The equation governing the behaviour of private consumption growth switches, however, when the floor on the short nominal interest rate becomes binding (when the economy is in a liquidity trap):

$$\dot{p} = b(c + g - \bar{y}) \quad (23)$$

$$\begin{aligned} \dot{c} &= [\bar{i} + (g - 1)p - d]c & \text{if } \bar{i} + gp \geq i_M \\ &= [i_M - p - d]c & \text{if } \bar{i} + gp < i_M \end{aligned} \quad (24)$$

When the liquidity trap constraint is not binding (we shall refer to this as the ‘normal’ case), saddlepoint stability for the dynamic system requires $g > 1$. A higher rate of inflation leads, through the policy reaction function, to a larger increase in the short nominal bond rate so as to raise the short real rate. As shown in Chart 1, the $\dot{c} = 0$ locus in the normal case (denoted

(8) See Buiter and Panigirtzoglou (1999) for a longer version of the paper which includes an analysis of the liquidity trap with flexible prices.

$(\dot{c} = 0)_N$) is vertical in a phase diagram with \mathbf{p} on the horizontal axis and c

on the vertical axis, at $\mathbf{p} = \frac{\bar{i} - \mathbf{d}}{1 - \mathbf{g}} = \mathbf{p}^*$.⁽⁹⁾

In the liquidity trap regime, the $\dot{c} = 0$ locus (denoted $(\dot{c} = 0)_L$) is vertical at $\mathbf{p} = i_M - \mathbf{d}$. With $i_M = 0$, the locus $(\dot{c} = 0)_L$ is to the left of $(\dot{c} = 0)_N$. This is the case we shall be considering as the benchmark henceforth.

As long as the rate of inflation exceeds $\frac{i_M - \bar{i}}{\mathbf{g}}$, the short nominal bond rate exceeds the nominal interest rate on currency, and the economy is in the normal regime. For inflation rates at or below $\frac{i_M - \bar{i}}{\mathbf{g}}$, the economy is in the

liquidity trap regime. The switch from the normal to the liquidity trap regime occurs at $\mathbf{p} = \frac{i_M - \bar{i}}{\mathbf{g}} = \frac{i_M - \mathbf{d}}{\mathbf{g}} + \left(\frac{\mathbf{g} - 1}{\mathbf{g}} \right) \mathbf{p}^*$. We shall refer to the

boundary of the normal and the liquidity trap locus as the LN locus in Chart 1. Taking again as our reference point the situation where $i_M = 0$, and noting that $\mathbf{d} > 0$ and that $\mathbf{g} > 1$, we need only assume that the target inflation rate \mathbf{p}^* is non-negative, for the switching value of \mathbf{p} to lie between the two $\dot{c} = 0$ loci. This is assumed in Chart 1 and thereafter. The LN locus could either be to the left or to the right of the c axis.

The steady state of the model is as follows. There are two steady states (the normal one and the liquidity trap case) for the nominal bond rate and the rate of inflation. The normal steady state values are given first.

$$c = \bar{y} - g$$

$$r = \mathbf{d}$$

(9) Here and in what follows we ignore the $c = 0$ segment of the c isocline.

$$\mathbf{p} = \frac{\mathbf{d} - \bar{i}}{\mathbf{g} - 1} = \mathbf{p}^* \quad (\text{Normal case})$$

or

$$\mathbf{p} = \bar{i}_M - \mathbf{d} \quad (\text{Liquidity trap})$$

$$i = \frac{\mathbf{g}\mathbf{d} - \bar{\mathbf{i}}}{\mathbf{g} - 1} \quad (\text{Normal case})$$

or

$$i = \bar{i}_M \quad (\text{Liquidity trap})$$

When $\mathbf{g} > 1$, as we assume throughout, the equilibrium configuration in the neighbourhood of the normal steady state is a saddlepoint.

The linear approximation of the normal dynamics at $c = \bar{c}$ and $\mathbf{p} = \bar{\mathbf{p}}$ is:

$$\begin{bmatrix} \dot{c} \\ \dot{\mathbf{p}} \end{bmatrix} \approx \begin{bmatrix} \bar{i} + (\mathbf{g} - 1)\bar{\mathbf{p}} - \mathbf{d} & (\mathbf{g} - 1)\bar{c} \\ \mathbf{b} & 0 \end{bmatrix} \begin{bmatrix} c - \bar{c} \\ \mathbf{p} - \bar{\mathbf{p}} \end{bmatrix}$$

At the normal steady state, with $\bar{c} = \bar{y} - \mathbf{g}$ and $\bar{\mathbf{p}} = \frac{\mathbf{d} - \bar{i}}{\mathbf{g} - 1}$, this reduces to:

$$\begin{bmatrix} \dot{c} \\ \dot{\mathbf{p}} \end{bmatrix} \approx \begin{bmatrix} 0 & (\mathbf{g} - 1)(\bar{y} - \mathbf{g}) \\ \mathbf{b} & 0 \end{bmatrix} \begin{bmatrix} c - \bar{c} \\ \mathbf{p} - \bar{\mathbf{p}} \end{bmatrix}$$

The determinant of the state matrix is $(1 - \mathbf{g})(\bar{y} - \mathbf{g})\mathbf{b} < 0$ if $\mathbf{g} > 1$. The two characteristic roots are $\pm\sqrt{\mathbf{b}(\mathbf{g} - 1)(\bar{y} - \mathbf{g})}$.

The normal steady-state configuration Ω^N in Chart 1 illustrates the saddlepoint property of the normal steady state. From (23) and the normal version of (24) it follows that the slope of the integral curves in $c - \mathbf{p}$ space is given by:

$$\frac{dc}{d\mathbf{p}} = \frac{[\bar{i} - \mathbf{d} + (\mathbf{g} - 1)\mathbf{p}]c}{\mathbf{b}(c + \mathbf{g} - \bar{y})}$$

This can be rewritten as:

$$\mathbf{b} \left(1 + \frac{g - \bar{y}}{c}\right) dc = [\bar{i} - \mathbf{d} + (g - 1)\mathbf{p}] d\mathbf{p}$$

As this is separable in c and \mathbf{p} , it can be integrated to yield

$$\mathbf{b}[c + (g - \bar{y})\ln c] = (\bar{i} - \mathbf{d})\mathbf{p} + \frac{(g - 1)}{2}\mathbf{p}^2 + k$$

where k is an arbitrary constant.

Provided $(\bar{i} - \mathbf{d})^2 + 2(1 - g)(k - \mathbf{b}[c + (g - \bar{y})\ln c]) \geq 0$, the integral

curves in the normal case ($c > 0$, $\mathbf{p} > \frac{i_M - \bar{i}}{g}$) are given by:

$$\mathbf{p} = \frac{\bar{i} - \mathbf{d} \pm \sqrt{(\bar{i} - \mathbf{d})^2 + 2(1 - g)(k - \mathbf{b}[c + (g - \bar{y})\ln c])}}{1 - g}$$

The equilibrium configuration near the liquidity trap steady state (Ω^L in Chart 1) is neutral and cyclical (the linearised dynamic system at Ω^L has two complex conjugate roots with zero real parts). The integral curves for the

liquidity trap case ($c > 0$, $\mathbf{p} \leq \frac{i_M - \bar{i}}{g}$) are given by

$$\mathbf{p} = i_M - \mathbf{d} \pm \sqrt{(i_M - \mathbf{d})^2 + 2(k - \mathbf{b}[c + (g - \bar{y})\ln c])}$$

The liquidity trap configuration is a *center*.⁽¹⁰⁾ Some neighbourhood of this steady state is completely filled by closed integral curves, each containing the steady state in its interior. The left-hand panel of Chart 1 shows the behaviour of the system near the liquidity trap steady state.

(10) Anne Sibert provided the mathematical solution and graphical representation for the behaviour of the system in the liquidity trap region.

At a common level of consumption, the slope of the integral curve in the

normal case, $\left. \frac{dc}{d\mathbf{p}} \right|_N$ is the same as the slope of the integral curves in the

liquidity trap case $\left. \frac{dc}{d\mathbf{p}} \right|_L$ on the boundary of the two regimes (when

$\mathbf{p} = \frac{i_M - \bar{i}}{g}$). It is easily checked that

$$\left. \frac{dc}{d\mathbf{p}} \right|_{\mathbf{p} = \frac{i_M - \bar{i}}{g}}^N = \left. \frac{dc}{d\mathbf{p}} \right|_{\mathbf{p} = \frac{i_M - \bar{i}}{g}}^L = \frac{\left[\left(\frac{g-1}{g} \right) i_M + \frac{1}{g} \bar{i} - d \right] c}{b(c + g - \bar{y})}$$

The orbits on the left-hand side of Chart 1 are derived on the assumption that

$i = i_M$ throughout. This means that not only the instantaneous short nominal interest rate is at its floor, but that the nominal interest rates at all other maturities, from zero to infinity, are also at that floor (i_M). Solution trajectories that switch between a liquidity trap orbit and a normal solution curve will not, in general, have all longer-maturity nominal yields equal to the floor value.

Note that, for any initial (and predetermined) rate of inflation that starts the

economy off in the liquidity trap region ($\mathbf{p} < \frac{i_m - \bar{i}}{g}$), there will be a

continuum of possible liquidity trap solutions. Each closed liquidity trap orbit that lies entirely to the left of the boundary between the liquidity trap region and the normal region will have two equilibrium values of consumption corresponding to any given rate of inflation. These solution orbits fill the entire plane. The model does not offer criteria for choosing among this continuum of possible solutions.

This means that, even when the system starts off with a predetermined initial rate of inflation above the value that separates the liquidity trap region from the normal region, if the exogenous variables generate a solution trajectory that must, at some point, leave the normal region and enter the liquidity trap region, there will be a continuum of solutions for the initial consumption

level.⁽¹¹⁾ Thus, if, with initial inflation above $\frac{i_m - \bar{i}}{g}$, there is no initial value of

consumption from which the economic system will not eventually enter the liquidity trap region, then there will also be a continuum of initial solution values for consumption.⁽¹²⁾

We want to consider shocks for which the liquidity trap can be sprung, ie shocks for which the constraint $i \geq i_M$ becomes binding. In our model, that has to be either a demand shock or a supply shock that lowers current aggregate demand below current capacity output. As regards demand shocks, the simplest candidates in our model are either the unexpected announcement, at time $t = t_0$, of a future permanent increase in public spending g , starting at $t_1 > t_0$, or the unexpected announcement, at time $t = t_0$, of an immediate temporary reduction in public spending. For reasons of space, we restrict the presentation to the anticipated future public spending increase.

An anticipated future increase in public spending is contractionary between the announcement date (t_0) and the implementation date (t_1) because forward-looking Ricardian households realise that higher future public spending means a higher present discounted value of future taxes. Human capital falls immediately, and private consumption falls with it, while the

(11) We maintain the assumption that explosively divergent solution trajectories are ruled out.

(12) We are indebted to Nobu Kiyotaki and Naoki Shimoi for convincing us of the need to clarify this issue.

increase in future public spending has not yet materialised. The consumption function for our model is:

$$c(t) = \frac{d}{1+h} \left[\frac{M(t) + B(t)}{P(t)} + \int_t^\infty e^{-\int_t^v [i(u) - p(u)] du} [y(v) - t(v)] dv \right] \quad (25)$$

With the logarithmic utility function, the intertemporal substitution elasticity is unity, and the marginal propensity to spend out of comprehensive wealth,

$\frac{d}{1+h}$, is independent of current and anticipated future real interest rates.

Real interest rates affect current consumption because they discount future real after-tax endowments. Monetary policy affects consumption to the extent that changes in current and anticipated future short nominal rates can affect current and anticipated future real discount factors,

$$\int_t^\infty e^{-\int_t^v [i(u) - p(u)] du}, \text{ at any horizon } v - t \geq 0.$$

Because of the Taylor-style interest rate reaction function, the profile of expected future short real rates is actually lower with the public spending shock than without. Future after-tax endowments are therefore discounted at a lower rate, but this is not enough to negate the negative effect on aggregate demand of the higher sequence of future taxes.⁽¹³⁾ Chart 2 represents the behaviour of the system following the public spending shock.

To understand the solution, and the multiplicity of solutions if the system is destined to enter the liquidity trap region, three properties of admissible solutions are key. First, explosively divergent solutions are ruled out, if non-explosive solutions exist. Second, the inflation rate is predetermined. Third, discontinuous jumps in the level of private consumption (the

(13) If instead of the logarithmic instantaneous utility function we had adopted the constant elasticity of marginal utility function with an intertemporal substitution elasticity larger than 1, the negative effect on consumption would have been reinforced.

non-predetermined state variable) are permitted only at instants that news arrives. In what follows, news arrives at the initial date.

For the time being, we will make the (arbitrary) assumption that if, starting off in the normal region, there exists a solution that stays within the normal region forever, this is the solution that will in fact be selected, even though there may be other solutions that exit from the normal into the liquidity trap region. We will refer to this as the *normal invisible hand assumption*, or *NIHA*.

Assume the system starts, before the news arrives, in steady state at the normal steady-state equilibrium Ω_1^N , with government spending expected to be constant. Given the *NIHA*, an unanticipated, immediate, permanent increase in public spending will result in an immediate transition to the new steady state at Ω_2^N . In the new steady state, the rate of inflation and all real and nominal interest rates are the same as before. The level of private consumption falls by the same amount as the increase in the level of public consumption.

When the increase in public consumption is not immediate, the transition is as follows. Assume that at the announcement date, t_0 , there is unexpected news of a future permanent increase in public spending, starting at $t_1 > t_0$. The increase in public spending, when it occurs, is of the same magnitude as the immediate increase in public spending analysed earlier. We shall refer to t_1 as the implementation date. For a ‘moderate’ postponement in the implementation of the public spending increase (defined as a postponement that still permits complete solution trajectories that do not leave the normal

region), private consumption drops immediately to a point like Ω_{12}^N , between Ω_1^N and Ω_2^N .⁽¹⁴⁾

The reason again is that a higher sequence of future taxes is anticipated by the Ricardian consumers, immediately upon the announcement of the future spending increase. Human capital falls and with it private consumption, albeit by less than when the public spending increase was immediate. Between the announcement date, t_0 , and the implementation date, t_1 , consumption and inflation both fall gradually as the system moves from Ω_{12}^N to Ω_{22}^N , where it arrives at t_1 when the public spending increase is actually implemented. From t_1 on, the system moves from Ω_{22}^N along the convergent saddlepath through the new steady state at Ω_2^N .

The initial jump in the level of consumption at t_0 is such as to place the system on that divergent trajectory, drawn with reference to the initial steady state, that will put it on the unique continuously convergent trajectory through the new steady state, Ω_2^N , at the moment the public spending increase is actually implemented (at the implementation date t_1). Note again that the rate of inflation is assumed to be predetermined in this Keynesian model, with its backward-looking inflation adjustment equation.

A *moderate* delay in the public spending increase is defined by the requirement that the intersection of the disequilibrium trajectory, drawn with reference to the initial steady state, and the saddlepath through the new steady state be at a level of inflation greater than or equal to the one that triggers the liquidity trap. In Chart 2 this means that the disequilibrium

(14) We again make the invisible normal hand assumption.

trajectory passing through Ω_{13}^N is the trajectory with the longest gap between the announcement of the future spending increase and its implementation that is consistent with the system not ending up in the liquidity trap region. This solution trajectory intersects the convergent saddlepath through Ω_2^N at Ω_{23}^{LN} , which corresponds to an inflation rate equal to $\frac{i_M - \bar{i}}{g}$, which separates the normal region from the liquidity trap region.

With any longer postponement of the public spending increase, the initial drop in consumption would be to some level between Ω_3^N and Ω_1^N , say, Ω_{14}^N in Chart 2. The *NIHA* now is inoperative, as there exists no non-explosively divergent solution that does not at some point enter the liquidity trap region. From Ω_{14}^N the system would travel along the divergent trajectory drawn with reference to Ω_1^N that would bring it to Ω_{24}^{LN} on the *LN* locus some time before t_1 . The system would then switch to the closed orbit, drawn with reference to Ω_1^L , that passes through Ω_{24}^{LN} . It would travel in clockwise fashion around this orbit until t_1 . Assume that at t_1 it has arrived at Ω_{34}^L . At t_1 , it would switch (without a discontinuous jump in either c or \mathbf{p}) to the closed orbit drawn with reference to the new liquidity trap steady state (Ω_2^L) that passes through Ω_{34}^L . If this orbit (labelled \mathbf{aw}) stays entirely within the liquidity trap domain (ie if this orbit does not cross the *LN* locus) the system would continue to circumnavigate the new liquidity trap steady state on this closed orbit. This is the case drawn in Chart 2. If the \mathbf{aw} orbit leaves the liquidity trap domain again, the behaviour of the system becomes very hard to pin down. It is possible that no equilibrium exists in this case.

Note that any initial value of consumption between Ω_3^N and Ω_1^N would be an equilibrium, once the postponement of the future public spending increase is sufficient to rule out a solution trajectory that lies entirely inside the normal region. More than that, if we drop the *NIHA* (which was of course quite arbitrary), there will be two kinds of solutions for any public spending postponement that do support solution trajectories that stay entirely within the normal region. The first is the (unique) solution that remains entirely within the normal region. The second is the continuum of solutions between Ω_3^N and Ω_1^N . For instance, in response to the unexpected announcement of an immediate permanent increase in public spending, private consumption could either fall to Ω_2^N and stay there, or fall to any level between Ω_3^N and Ω_1^N , and move from there into the liquidity trap region.

If the *NIHA* applies, it is clear how a reduction in the exogenous nominal interest rate on currency, i_M , can help to avoid a liquidity trap. A lower value of i_M shifts both the $(\dot{c} = 0)_L$ locus and the boundary separating the liquidity trap region from the normal region (the *LL* locus) to the left. For any shock to demand or supply, it is always possible to find a value of i_M low enough to stop the economy from entering the liquidity trap region. Furthermore, even if the *NIHA* does not apply, if the economy were to get caught in the liquidity trap region, an unexpected permanent reduction in the nominal interest rate on currency could always place the predetermined rate of inflation back in the normal region. However, unless the *NIHA* applies, having the initial inflation rate back inside the normal region does not guarantee that the non-predetermined consumption level will, following the (unexpected) reduction in i_M , take on a value that will keep the solution trajectory inside the normal region, even if such a solution exists. If the *NIHA* applies, cutting the nominal rate on currency can therefore be cure as well as

prevention.⁽¹⁵⁾ It will become clear in Section 4 that, though the algebra of negative nominal interest rates on currency is trivial, the practical implementation of such a policy is likely to be very cumbersome.

3. Can the zero nominal interest rate floor become binding in the United Kingdom?

Most estimates of the current level of the long real interest rate in the United Kingdom put it somewhere between 2.0% and 3.0% per annum. Charts 3 and 4 show the recent behaviour of medium-term and long-term real rates of interest on index-linked government securities.⁽¹⁶⁾

(15) A very similar, but technically more complicated, analysis can be conducted for the case where the growth rate of the nominal money stock rather than the short nominal bond rate is the monetary instrument. When the growth rate of nominal money is the exogenous monetary instrument, the equations of motion of the economic system can be summarised as follows: When $i > i_M$, we have the following three-dimensional dynamic system:

$$\frac{d}{dt} \left(\frac{c}{m} \right) = (i_M + \mathbf{h} \frac{c}{m} - \mathbf{d} - \mathbf{m}) \frac{c}{m}$$

$$\dot{c} = (i_M + \mathbf{h} \frac{c}{m} - \mathbf{d} - \mathbf{p})c$$

$$\dot{\mathbf{p}} = \mathbf{b}(c + g - \bar{y})$$

Note that when $i = i_M$ and the economy is stuck in a liquidity trap, the dynamic system reduces to:

$$\dot{c} = (i_M - \mathbf{d} - \mathbf{p})c$$

$$\dot{\mathbf{p}} = \mathbf{b}(c + g - \bar{y})$$

which is the same as the liquidity trap-constrained dynamics when the short nominal bond rate was the policy instrument.

(16) Jenny Salvage prepared Charts 3 to 7.

With an inflation target of 2.5% per annum (as in the United Kingdom), the long-run nominal interest rate (ignoring term and risk premia) would be between 4.5% and 5.5% per annum. In steady state, the short-term nominal interest rate would also be between 4.5% and 5.5% per annum. We can regard this as the ‘normal’ level of the short nominal interest rate. If one believed that there were contingencies (such as a dramatic, spontaneous collapse of aggregate demand) under which a cut in short rates of more than 4.5% to 5.5% would be in order, the monetary authority would be at risk of hitting the zero interest floor.

Historically, in the United Kingdom, there have been occasions when Bank Rate has swung by more than 4.5 or 5.5 percentage points. On 15 November 1979, the Bank’s Minimum Lending Rate hit 17%. On 11 March 1981, it stood at 12%. On 6 October 1989, the Bank’s Minimum Band 1 Dealing Rate stood at 14.88%. On 8 February 1994, it was down to 5.13%. Clearly, very large swings in Bank Rate, in excess of the 4.5% or 5.5% ‘safety margin’ associated with a 2.5% inflation target and a 2.0% to 3.0% long real interest rate, have occurred in the past.

The emphasis should, however, be on ‘in the past’. These very large cuts in Bank Rate invariably took place from a very high level of rates associated with prior macroeconomic mismanagement, generally an inflationary surge that threatened to get out of control (or had indeed done so) or the desperate defence of an overvalued exchange rate peg. Chart 5 shows the behaviour of Bank Rate (or its successor rates), the inflation rate and the sterling-US dollar exchange rate for the United Kingdom in the post-World War II period.

Neither situation applies today. Nor should it apply again if the political commitment to low and stable inflation and its institutional expression in an operationally independent central bank remain intact (see also Johnson, Small and Tryon (1999) for a US perspective on this and related issues).

The longer-term historical record can also be viewed as encouraging. The United Kingdom got through the period 1800-1914 without ever landing itself in a liquidity trap. As Chart 6 shows, the average rate of inflation over this

115-year period was slightly negative and the variability of the inflation rate was high. Chart 7 shows that Bank Rate did not fall below 2% throughout the 115 years preceding World War I.⁽¹⁷⁾

There is a marked positive association, over time and across countries, between the level of the inflation rate and its variability (see Okun (1971, 1975), Taylor (1981), Ball and Cecchetti (1990)). If such a relationship were to be found also between the level of short nominal rates and their variability or volatility, it would further reduce the likelihood of ending up in a liquidity trap in an environment with sustained low inflation and therefore, on average, with low nominal interest rates.

As will become apparent, the available statistical evidence on the association between the level and volatility of short-term nominal interest rates is mixed, and neither weakens nor reinforces our prior belief, that it is hard to conceive of situations in which the zero nominal interest rate floor would become a binding constraint on monetary policy in the United Kingdom, with the current symmetric annual inflation target of 2.5%.

In principle, this hypothesis can be tested by estimating a dynamic stochastic process for the short nominal interest rate, using either time series or Markov chain models. By making distributional assumptions about the disturbances in this process, it would be possible to calculate the odds on the short nominal rate falling below zero, either conditionally, ie given the starting values of the process, or unconditionally. We do indeed attempt this, but our efforts must be accompanied by a clear health warning.

There is an obvious, and in our view virtually insurmountable, problem with any assessment, based on historical data, of the odds that the non-negativity constraint on short nominal rates will become binding. During the sample,

(17) The temporary collapse in the external value of the US dollar starting in 1861 reflects the exceptional circumstances of the American Civil War and its aftermath, the Greenback period.

markets undoubtedly were operating under the assumption that short nominal rates could never fall below zero. In the United Kingdom over the past 200 years, the annual Bank Rate series indeed never fell below 2%. With the support of the empirical distribution of nominal short rates truncated from below at zero, the historical interest rate record is unlikely to be informative about the odds on the economy getting into a liquidity trap in the future, since this would require a structural break in the interest rate process, about which the sample is uninformative. If we were to assume (counterfactually, as can be seen from Charts 9, 10 and 11) that the distribution of Bank Rate or of the error term in the Bank Rate equation is normal, there will always be a positive probability that Bank Rate will go negative. If we assume instead that the distributions in question are, say, lognormal, the probability of breaching the zero floor (even asymptotically) is *a priori* constrained to be zero. We try to circumvent this by calculating the asymptotic confidence bands for Bank Rate reported below from the empirical distribution of the sample residuals. Since the empirical distribution of the residuals obviously has finite support, this procedure will, if anything, underestimate the likelihood of the economy ending up in a liquidity trap.

Even ignoring the unavoidable small-sample problems, this procedure is vulnerable to the following criticism. What we are interested in is the probability that the short nominal interest rate would have had to be negative in order to avoid the economy getting into a very undesirable equilibrium⁽¹⁸⁾ If the economy had been in a liquidity trap in the sample, the data would reflect the liquidity trap configuration of the economy, including the endogenous responses of real activity and inflation that supported the liquidity trap floor as an equilibrium. Information on the ‘deep’ structural parameters of the model (the invariant parameters governing money demand and its determinants) is necessary to recover the ‘first passage’ probabilities into the liquidity trap region of the economy, and these cannot be recovered

(18) In principle, there could be a positive lower bound on the nominal interest rate, even with the own rate on currency at zero.

from the atheoretical, reduced-form time series processes we estimate. Finally, our statistical tests only concern the likelihood of the short nominal interest rate hitting the zero bound. As was pointed out in the introduction, this is only a necessary condition for the economy being in a liquidity trap. What do the data tell us about the statistical association between the level and volatility of the short nominal interest rate? The very high-frequency association between short nominal sterling rates and a measure of volatility derived from short sterling futures over the period 1987-99 is shown in Chart 8.⁽¹⁹⁾ The association between the level of short sterling and its volatility is, if anything, weakly negative.

The slightly lower frequency time series evidence on the association between the level of short nominal interest rate and a statistical measure of its variability using weekly data is also mixed. Chart 9 shows the time series record for the United Kingdom for the period 1997-99.

For the whole period 1975-99, volatility and level of the three-month interbank rate are positively contemporaneously correlated, but for the post-inflation targeting period 1993-99, the correlation is slightly negative. The statistical model that generated the conditional variance measure used in Chart 9 can be found in the Appendix. We also provide an estimate of the steady-state (long-run) value of the three-month interbank rate implied by the statistical model, together with 95% steady-state confidence bands for the three-month interbank rate.

We also investigated the statistical properties of Bank Rate at significantly lower frequencies, using a 200-year time series of annual observations. Our time series model (described in the Appendix) implies a strong positive correlation (0.81) between the level of Base Rate and its contemporaneous conditional variance. Chart 10 plots the level of Base Rate and our estimates

(19) Backing a volatility estimate out of futures prices is attractive, because it avoids the need to construct statistical estimates of volatility from the time series data on interest rates.

of its conditional variance. We also provide an estimate of the steady-state (long-run) value of Base Rate implied by the statistical model, together with 95% steady-state confidence bands for Base Rate.

The confidence bands were calculated using the distribution of the estimated sample residuals. Not surprisingly, the distribution of sample residuals is distinctly non-normal. The same holds for Base Rate itself. Chart 11 shows the frequency distribution of Base Rate and Charts 12 and 13 those of the estimated interest rate residuals. The sample distribution of Bank Rate is significantly skewed to the right. Its empirical distribution is truncated from below at 2.0%. The distribution of the sample residuals from the two main interest rate models is rather more symmetric.

Krugman (1998d) has suggested that deflation (negative inflation) makes a liquidity trap more likely. This is indeed an implication of just about any model of liquidity traps, including the model we developed in Section 2 of this paper. We therefore estimate a simple time series process for annual RPI inflation over the 200-year period, and for its conditional variance. The results are reported in Chart 14, together with its estimated steady-state value and steady-state 95% confidence intervals. The statistical inflation model is described in the Appendix. Surprisingly, the contemporaneous correlation between inflation and its conditional variance turns out to be negative.

The steady-state confidence intervals for the annual rate of RPI inflation show that there is quite a large probability of deflation. Before one gets too worried about this, three points should be kept in mind. First, the relationship between interest rates and expected inflation depends on the behaviour of the inflation risk premium. Second, the United Kingdom experienced negative trend inflation and short bouts of sharp deflation in the 19th century, without landing itself in a liquidity trap. Third, the monetary policy target in the United Kingdom is, since June 1997, a symmetric inflation target. Deviations of inflation below the 2.5% target are to be avoided as much as deviations above that target. The risk of sharp deflation

is therefore diminished. The new monetary regime has been in operation for too short a period, however, for this to show up as a structural break in the inflation time series process.

McKinnon and Ohno (1999) have argued that, at any rate in the Japanese case, a large expected appreciation of the currency could create a liquidity trap. We investigated the likelihood of a sharp appreciation of sterling by using almost 200 years of £/\$ exchange rate data to estimate a simple stochastic process for the proportional rate of depreciation of the exchange rate and its conditional variance. The results are reported in Chart 15, together with the expected long-run sterling depreciation rate and the 95% asymptotic confidence intervals. The statistical model underlying these calculations is described in the Appendix. The contemporaneous correlation between exchange rate depreciation and its conditional variance is low but negative.

It suggests that, based on this particular statistical model, there is quite a significant probability of a sizeable appreciation of sterling. Again, the caveat about the dangers of ignoring risk premia applies. It is surprising that our simple statistical model appears to handle such episodes as the American Civil War, two World Wars and the Great Depression of the 1930s quite well.

Finally, Chart 16 reports cross-sectional evidence on the relationship between the level of short nominal rates and their volatility based on a sample of 59 countries between 1989 and 1998. The source of the data is IFS.⁽²⁰⁾ The correlation between the two variables is very high at 0.89, suggesting that across countries high short-term nominal rates are accompanied by high unconditional variances.

(20) We would like to thank Nick Hanchard for preparing this chart.

Orphanides and Wieland (1998) use stochastic simulations of a small structural rational expectations model calibrated for the United States, to investigate the consequences of the zero bound on nominal interest rates. They found that if the economy is subject to stochastic shocks similar in magnitude to those experienced by the United States over the 1980s and 1990s, the consequences of the zero bound were negligible for target inflation rates as low as 2%. With target inflation between 0% and 1%, there was a quantitatively significant deterioration in economic performance.

On balance, the data fail to offer convincing support either for or against the contention that a regime of low short nominal interest rates is likely to be a regime of stable short nominal interest rates. This is therefore not unambiguously good news, nor unambiguously bad news for a policy-maker targeting low inflation, although the current UK target would seem to provide quite handsome room for monetary manoeuvre. Two hundred years of UK monetary history also favour the contention that zero bounds on short nominal interest rates are unlikely to become a policy concern.

4. Options for avoiding a liquidity trap

The lower the inflation target (if it is credible) and the lower the underlying rate of inflation, the narrower is the 'safe range' above the zero floor for the short nominal rate. A credible target of zero inflation, would, with the long real rate at 2.0% to 3.0%, reduce the safe range to 2.0% to 3.0%.

Does this mean that targeting zero inflation would be a high-risk strategy? There would be risks if, despite a credible commitment to zero inflation, the economy were likely to be hit by shocks that would make interest rate cuts of more than 200 or 300 basis points desirable. In the worst-case scenario, the economy could end up in a liquidity trap with monetary policy incapable of influencing any variable, nominal or real, that matters. If fiscal policy cannot

be used to escape from the liquidity trap and if the risk of ending up in a trap is considered unacceptable, two options remain.

Raising the inflation target

The first option is to accept the nominal interest rate floor as immutable, and to target a rate of inflation high enough to reduce to acceptable levels the risk of hitting the zero interest floor. This would have to be done before the country gets into a liquidity trap. Targeting a higher rate of inflation once you are caught in the trap would be pointless.⁽²¹⁾ With a given nominal interest rate floor, only an exogenous shock to excess demand in the output market (for instance, one administered through an expansionary fiscal policy measure), can boost the economy out of the trap.

Lowering the nominal interest rate floor: stamping money à la Gesell

The only other option is to stick to the inflation target but to lower the floor on the nominal interest rate. Setting the floor at a level below zero would reduce the likelihood that the floor would ever become a binding constraint on policy. In addition, the option of further lowering the floor would provide a mechanism for getting off the floor, even after the constraint had become binding.

That nominal interest rate floor at zero is not a technological, immovable barrier. It is the result of a policy choice – the decision by governments or central banks to set the administered nominal interest rate on coin and currency at zero, rather than at some other (negative) level. Coin and

(21) Note again that a liquidity trap involves more than just the short nominal rate at its lower bound. In the model of Section 2, nominal rates at all maturities have to be at their lower (zero) bounds for the economy to be caught in a liquidity trap. In more general models, rates of return on all assets, private and public, short and long, financial and real, domestic and foreign currency denominated, have to be immune to changes in the quantity of money.

currency are government *bearer* bonds.⁽²²⁾ A bearer bond is a debt security in paper form whose ownership is transferred by delivery rather than by written notice and amendment to the register of ownership. We shall refer to all securities that are not bearer bonds as *registered* securities. Bearer bonds

(22) Bearer securities are securities for which ownership is established by possession, without any need for registering title. Thus, a bearer bond is a bond with no owner information attached to it. The legal presumption is that the bearer is the owner. If the issuer of the bond is credit-worthy, they are almost as liquid and transferable as cash. Cash (coin and currency) is a special case of a zero interest (or zero-coupon) bearer bond issued by the state (generally through the central bank). Currency can be viewed as a zero coupon bearer consol or bearer perpetuity, since it can be interpreted as having an infinite maturity. It may actually be more informative to view currency as a zero coupon finite maturity bearer bond, which is issued and redeemed at par, with redemption taking the form of the one-for-one exchange of old currency for new currency which is indistinguishable from the old currency (see Buiter and Panigirtzoglou (1999, Appendix 1)).

The vast majority of 'international bonds', historically called 'eurobonds' are bearer. Bearer bonds can take two main forms. First, the traditional 'definitive' style, where the bonds literally are individual pieces of security-printed paper in denominations of, say, \$10,000, which individual holders bring in to paying agents so as to receive payment of interest and principals. Second, 'global' bonds, which are technically bearer instruments but consist of a single piece of paper representing the entire issue (and so worth hundreds of millions or even billions of dollars). In practice, the terms of the global bond say that only Euroclear (the settlement system based in Brussels) or Cedelbank (the settlement system based in Luxembourg) are entitled to the proceeds of the global bond, and that Euroclear and Cedelbank will in turn divide the proceeds up among the end-investors whose details are stored in their electronic records. Thus the global bond is not an instrument which in practice can be passed from one owner to another, even though it is technically 'bearer'. Effectively the bonds are dematerialised.

Bearer bonds are legal and quite common in the United Kingdom. While the bearer debenture went out of use, replaced by the non-negotiable debenture or debenture stock, transferable (in the same way as common stocks) by entry in the company's register, a number of new negotiable investment securities have evolved. They include the modern bearer bond, the negotiable certificate of deposit, and the floating rate note. A limited number of gilts have also been issued with a bearer option.

Before July 1983, municipal securities in the United States were issued for the most part in certificate form with coupons attached. Some of these so-called old-style bearer bonds are still available in the marketplace. The issuer has no record of who owns these bonds. The owner clips the coupons and collects the interest from the issuer's paying agent. Transferring the bonds requires physical delivery and payment. Bearer bonds issued by municipal authorities were made illegal in the United States in 1982.

are negotiable, just as money market instruments such as Treasury bills, bank certificates of deposit, and bills of exchange, for example, are negotiable.⁽²³⁾ Coin and currency therefore are bearer bonds. They are obligations of the government, made payable not to a named individual or other legal entity, but to whoever happens to present it for payment – the bearer. Coin and currency have three further distinguishing properties: they are government bearer bonds with infinite maturities (perpetuities or consols),⁽²⁴⁾ their coupon payments (which define the own (or nominal) rate of interest on coin and currency) are zero, and they are legal tender (they cannot be refused in final settlement of any obligation).

There are two reasons why interest is not paid on currency.⁽²⁵⁾ The first and currently less important one has to do with the attractions of seigniorage (issuing non interest bearing monetary liabilities) as a source of government revenue in a historical environment of positive short nominal rates on non-monetary government debt.⁽²⁶⁾

The second, and more important, reason why no interest is paid on coin and currency is the practical, administrative difficulties of paying a negative interest rate on bearer bonds. Significant ‘shoe leather’ costs are involved both for the state and for private agents.

(23) A financial instrument is negotiable if it is transferable from one person to another by being delivered with or without endorsement so that the title passes to the transferee. Key elements of negotiability include the following: (1) transfer by physical delivery; (2) transfer is such as to confer upon its holder unchallengeable title and (3) a negotiable instrument benefits from a number of evidential and procedural advantages in the event of a court action.

(24) But see footnote 17 and Buiters and Panigirtzoglou (1999, Appendix 1) for a different interpretation.

(25) From here on, ‘currency’ will be taken to include both coin and currency. There obviously are more severe technical problems with attaching coupons or stamps to coin than to currency notes.

(26) Of course, issuing negative interest-bearing monetary liabilities would be even more attractive, from a seigniorage point of view.

There is no practical or administrative barrier to paying negative nominal interest rates (market-determined or administered) on registered securities, including balances held in registered accounts, such as bank accounts.⁽²⁷⁾ The reason is that, for registered securities, the identities of both the issuer and the holder (the debtor and the creditor) are known or easily established. This makes it easy to verify whether interest due has been paid and received. Thus the non-bearer bond part of the monetary base, ie banks' balances with the central bank, could earn a negative nominal interest rate without any technical problems. Positive interest payments or negative interest payments just involve simple bookkeeping transactions, debit or credit, between known parties.

There are technical, administrative problems with paying negative interest on the bearer bond part of the central bank's monetary liabilities, coin and currency. While the identity of the issuer (the central bank) is easily verified, the identity of the holder is not. There is no obligation to register title to currency in order to establish ownership. Possession effectively provides complete title. This creates problems for paying any non-zero interest rate, because it is difficult to verify whether a particular note or coin has already been credited or debited with interest.

The problem of verifying whether interest due on bearer bonds has been paid is present even when the interest rate is positive. However, the problem of getting the anonymous holder of currency to come forward to claim his positive coupon receipt from the government is less acute than the problem of getting the anonymous holder to come forward to make a payment to the government.⁽²⁸⁾ In both cases, however, each individual currency claim has to

(27) The only exception is that it would not be possible to have a consol or perpetuity with a negative nominal interest rate. Assume the constant nominal coupon payment of the consol is positive. If the infinite sequence of short nominal rates is negative, the value of the consol would be unbounded positive. A negative coupon would yield an unbounded negative value for the consol.

(28) This is akin to the problem of compelling payment of taxes when the tax base cannot be verified.

be marked clearly as 'current', ie as having paid or received all interest that is due. Without this, positive interest-bearing currency could be presented repeatedly for the payment of interest. Historically, the problem of paying positive interest on bearer bonds was solved by attaching coupons or stamps to the title certificate of the bearer bond. When claiming his periodic coupon payment, the appropriate coupon was physically removed ('clipped') from the title certificate and retained by the issuer.

Without further amendment, the 'coupon clipping' or stamping route would not work for bearer bonds with negative coupons. The enforcement problems involved in getting the unregistered, anonymous holders of the negative coupon bearer bonds to come forward to pay the issuer would be insurmountable. The only practical way around this problem is to make the bearer bond subject to an expiration date and a conversion procedure. In the case of currency, this could be achieved by periodically attaching coupons or stamps to currency, without which the currency would cease to be 'current'.

For currency to cease to be 'current', it is not enough for the monetary authority to declare that after a certain date 'old' currency shall cease to be legal tender. Being legal tender certainly enhances the attractiveness of currency as a store of value, medium of exchange and means of payment, but these advantages need not be enough to induce holders of 'old' currency, which is about to lose its legal tender status, to come forward and exchange it, at a price, for 'new' currency which does have continuing legal tender status. What serves as a medium of exchange and means of payment is socially determined. Being legal tender is but one among many considerations that induce people to use certain classes of object as a means of payment and medium of exchange. For currency to cease to be current, the bearer has to be subject to a serious penalty, such as confiscation, if the appropriate coupon or stamp has not been attached. In other words, there have to be periodic 'monetary reforms'.

There is a long tradition on the cranky fringes of the economics profession of proposals for taxing money or taxing liquidity. Many of these proposals were part of wide-ranging, and generally hare-brained, schemes for curing the world's economic and social ills. The mechanics of taxing currency are straightforward mainstream economics, however.

The best-known proponent of taxing currency was probably Silvio Gesell (1862-1930), a German/Argentinean businessman and economist admired by Keynes, who wrote of him 'I believe that the future will learn more from the spirit of Gesell than from that of Marx' (Keynes (1936, page 355)). Gesell wanted to stimulate the circulation of money by getting the state to issue money that, like capital assets, depreciated in value.⁽²⁹⁾ Rather than relying on inflation to reduce the attractiveness of holding money, Gesell proposed 'Stamp Scrip' - dated bills that would lose a certain percentage of value each year unless new stamps were put on them. Irving Fisher (1933) for a while supported the issuance of stamp scrip and wrote a sympathetic account of it. Stamp Scrip was actually issued briefly during the Great Depression of the 1930s in parts of the Canadian province of Alberta by the Social Credit provincial government of the day.⁽³⁰⁾ The Canadian federal government and the courts blocked the key measures, and in the end the provincial

(29) Gesell's motivation was not, as far as we can determine, the avoidance of or escape from liquidity traps. His aim was to eliminate the interest component of costs and prices completely from the economic system, not just in the extreme circumstances of the liquidity trap, but as a permanent feature. Our reading of his works suggest that he was a bit vague about the distinction between real and nominal interest rates. The formal model analysed in Section 2 of this paper has the property that the monetary authorities cannot influence the long-run real interest rate.

(30) In August 1935 the first social credit government was elected in the Canadian province of Alberta. While its ideology owed more to the writings of two other great economic cranks, Alfred Richard Orage (1917) and Major Clifford Hugh Douglas (1919) (and to the personal involvement of the latter as economic adviser to the provincial government), the Alberta Prosperity Certificates introduced in 1936 by Premier William Aberhart, were pure Gesell. Similar in appearance to a dollar bill, the certificates required a weekly endorsement of a 2c stamp, amounting to a 104% annual capital levy (see Hutchinson and Burkitt (1997) and Mallory (1954)).

government refused to accept its own scrip in payment.⁽³¹⁾ Similar local currency experiments were tried in Wörgl, Austria during the 1930s.

Thus, for negative interest on bearer bonds such as currency to be enforceable, the bearer bond has to expire after a certain date, unless it is converted into new currency. The desired interest rate on currency would be determined by the terms on which the old currency could be exchanged with the central bank for new currency. Taxing currency (or paying negative interest on currency) through expiration of old currency and conversion into new currency can be visualised as follows. After the expiration date, t_1 , the issuer (the central bank) or its agents can confiscate the old currency without compensation.⁽³²⁾ Provided the forces of the law are strong enough, this could induce holders of the old currency to convert it, at a price, on or before the expiration date, rather than continue to use it in transactions or as a store of value after the expiration date and risk having it confiscated. At fixed intervals of length Δt (Gesell periods, say) whose duration could, for convenience, be set at a year (or several years, in order to reduce conversion costs), and on a specific day, (Gesell day), old currency would legally revert to the issuer (the central bank). After Gesell day, the old currency has no value (because of the credible threat of confiscation) and will not be used in transactions or as a store of value. On Gesell day, £1 worth of new currency

(31) It also had failed to convince the Federal government in Ottawa to match its negative interest rates. Since Federal currency was at least as useful as a means of payment, this would require to scrip to trade at a discount with respect to the Federal currency and to appreciate *vis-à-vis* the Federal currency at a rate that compensated for the interest differential between Federal and provincial currency.

(32) Less drastic penalties might work also. For instance, old money found in circulation after its 'expiry' date would be forcibly converted into new money at the rate offered on the conversion date, but subject to an additional penalty. The confiscation scenario makes the key point very clearly, however.

would be issued in exchange for $e^{-i_M \Delta t}$ sterling worth of old currency, where i_M would be the policy-determined (instantaneous) nominal interest rate on currency.⁽³³⁾ For simplicity, we assume i_M to be constant, although it could be time-varying. The nominal rate of interest on currency would be administratively determined, ie set by the central bank. Earlier exchanges of old for new money might be allowed at the rate of £1 worth of the new

currency for $e^{-i_M \Delta t} e^{-\int_{t_e}^{t_1} i(s) ds}$ sterling worth of the old currency, where t_1 is the date of the next Gesell day, $t_e \leq t_1$ is the time before the next Gesell day on which the old currency is exchanged for the new, and i is the instantaneous nominal interest rate on the government's non-monetary liabilities. For currency to remain rate-of-return-dominated as a store of value, it is necessary that $i_M < i$. Both rates could be negative, and may have to be, if zero bounds are to be ruled out. Coin and currency would effectively become time-limited, finite-maturity financial claims.

New currency could, in principle, be used in transactions before midnight on the Gesell day before they are formally introduced. The relative value of the old currency in terms of the new currency would change at an instantaneous rate i_M , to ensure that, at the moment the old currency expires and the new currency comes in officially, there is no discrete jump in the value of old money in terms of new money, or of goods and services in terms of money.⁽³⁴⁾ It follows that, during the period of coexistence of old and new money, the rate of inflation of the prices of goods and services would be higher in terms of old money than in terms of new money, with the excess of the old money inflation rate over the new money inflation rate equal to $-i_M$.

(33) $e^{-i_M \Delta t} - 1$ would be the effective (Gesell) period tax rate on currency. The instantaneous tax rate would be $-i_M$.

(34) This is just like the ex-dividend price of a share of common stock being equal, on the day the dividend is paid, to the dividend-inclusive price of the stock minus the dividend. In our example, the dividend would be negative.

Our scheme for removing the zero nominal interest rate floor by taxing currency only applies to government bearer bonds with an administratively determined nominal rate of return, ie to coin and currency. Commercial banks' balances with the central bank are not bearer bonds, but registered securities, in the terminology of this paper. The nominal interest rate on these balances is determined administratively, but paying negative interest on them is as simple as paying positive interest. Bank deposits, which are private registered securities in our terminology, would not need to be taxed. If, when currency is taxed, the equilibrium nominal market yield on deposits, and on any other private registered securities, is negative, banks will pay a negative interest rate on deposits, without any need for taxing deposits. The same applies to private electronic or e-money, including 'money on a chip', Internet accounts etc.

Clearly there are costs associated with Gesell money, even if one can come up with a slightly higher-tech (and tamper-proof) alternative to physically stamping currency. These shoe leather costs have to be set against the benefits of removing the zero floor on the nominal interest rate.

There are costs (and benefits) other than shoe-leather costs associated with taxing currency. Taxing currency would be regressive, since only the relatively poor hold a significant fraction of their wealth in currency. Taxing currency would also, however, constitute a tax on the grey, black and outright criminal economies, which are heavily cash-based. In the case of the US dollar, with most US currency held abroad (one assumes by non-US residents), it would represent a means of increasing external seigniorage.⁽³⁵⁾

(35) Unless drug dealers switch elastically to non-stamped currency.

5. Conclusions

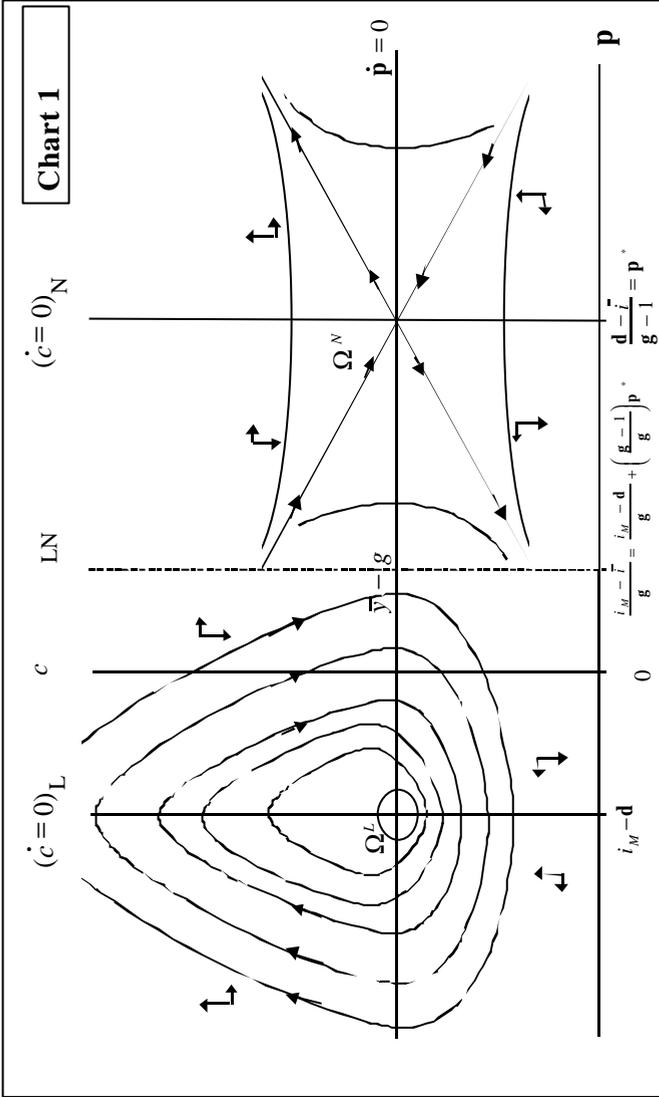
The credible targeting of a low rate of inflation should result, on average, in low nominal interest rates. The administratively determined zero nominal interest rate on currency sets a floor under the nominal interest rate on non-monetary financial claims. An important policy issue then is the following: how likely is it that the economy ends up, as a result of shocks or endogenous fluctuations, in a situation where the zero short nominal interest floor becomes a binding constraint? A short nominal interest rate at its lower floor is a necessary, although not a sufficient, condition for the economy to be in a liquidity trap, in which current and anticipated future monetary policy fails to have any impact on the economy. If low average nominal interest rates also tend to be stable rates, the risk of ending up in a liquidity trap need not be enhanced much by targeting a low rate of inflation. The empirical evidence on the relationship between the level and volatility of short nominal rates is, however, mixed. The cross-sectional evidence supports a strong positive correlation. The time-series evidence for the United Kingdom is ambiguous.

Once an economy lands itself in a liquidity trap, there are just two policy options. The first is to wait for some positive shock to the excess demand for goods and services, brought about through expansionary fiscal measures or through exogenous shocks to private domestic demand or to world demand. The second option is to lower the zero nominal interest rate floor on currency by taxing currency. A negative interest rate on currency would also reduce the likelihood of an economy landing in a liquidity trap.

The transactions and administrative costs associated with what amounts to periodic currency reforms would be non-trivial. Such currency conversion

costs could be reduced by lengthening the interval between conversions, but they would remain significant. These ‘shoe-leather costs’ of taxing currency have to be set against two kinds of potential benefits. If a low rate of inflation is targeted, taxing currency reduces the likelihood of ending up in a liquidity trap. If the alternative is a higher target rate of inflation, taxing currency avoids the distortions of a higher inflation tax rate.⁽³⁶⁾ It may take quite a lot of shoe leather to fill an output gap or to dominate a discounted sequence of Harberger triangles.

(36) On the costs of even low rates of inflation see Feldstein (1997), Tödter and Ziebarth (1997) and Chadha, Haldane and Janssen (1998). On the costs and benefits of low inflation see Akerlof, Dickens and Perry (1996). For a general survey see Fischer (1994).



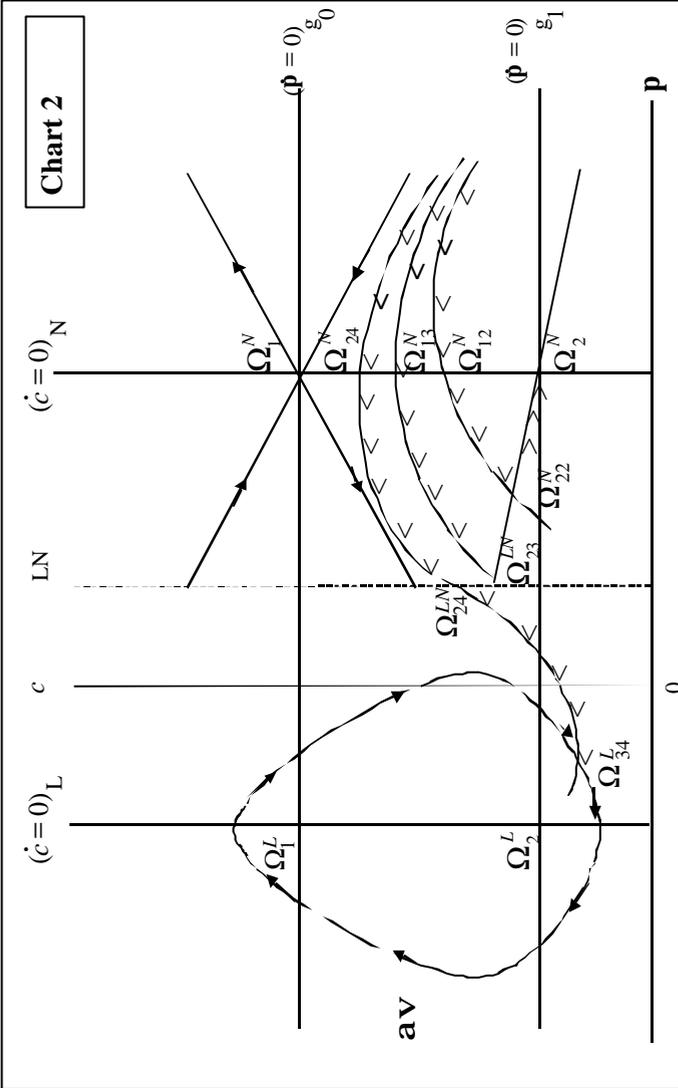
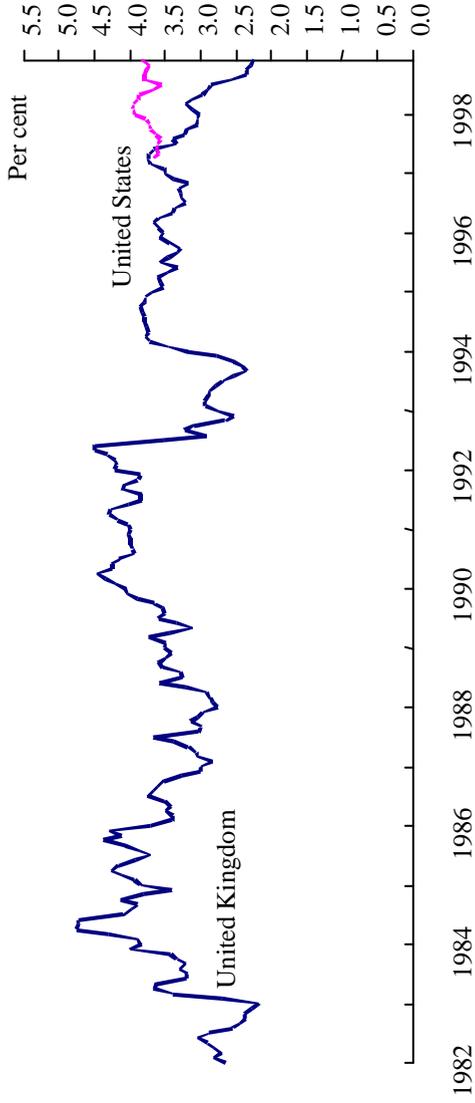


Chart 3

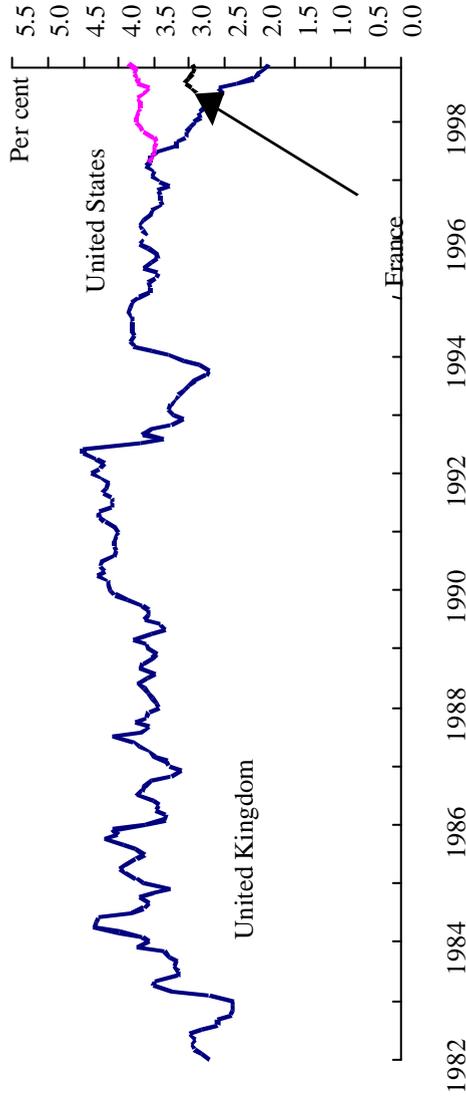
Five-year real interest rates in the United Kingdom and United States ^(a)



(a) US bonds mature in 2002.

Chart 4

Long term real interest rates in the United Kingdom, United States and France ^(a)



(a) UK bonds mature in 2009, US bonds in 2008, French bonds in 2009.

Chart 5
Bank Rate, inflation and £/\$ exchange rate, 1945 to date

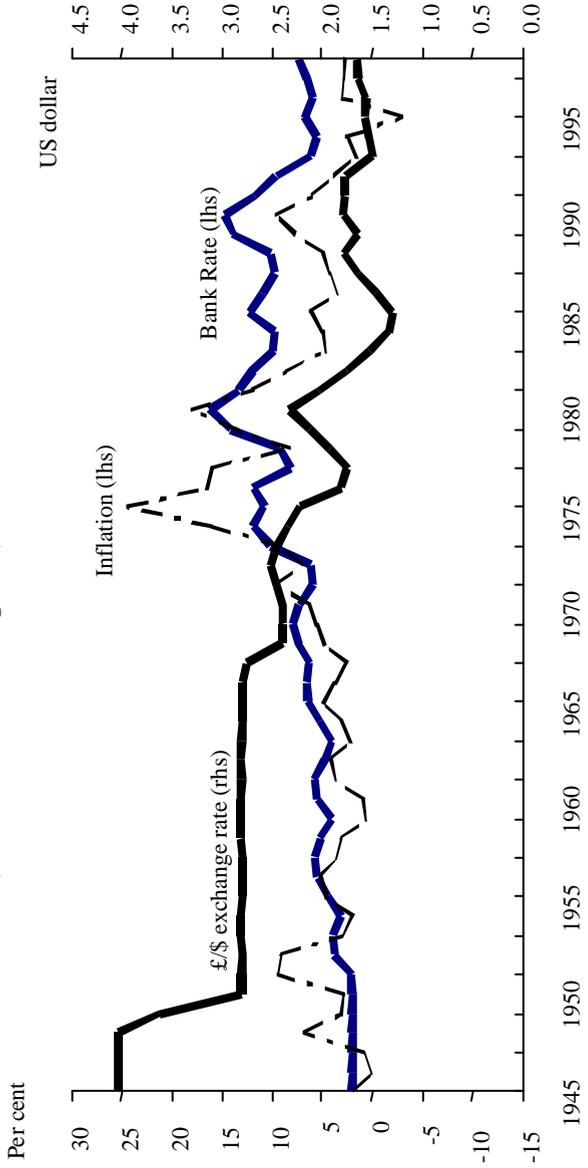


Chart 6
Price level and inflation in the United Kingdom, 1800-1914

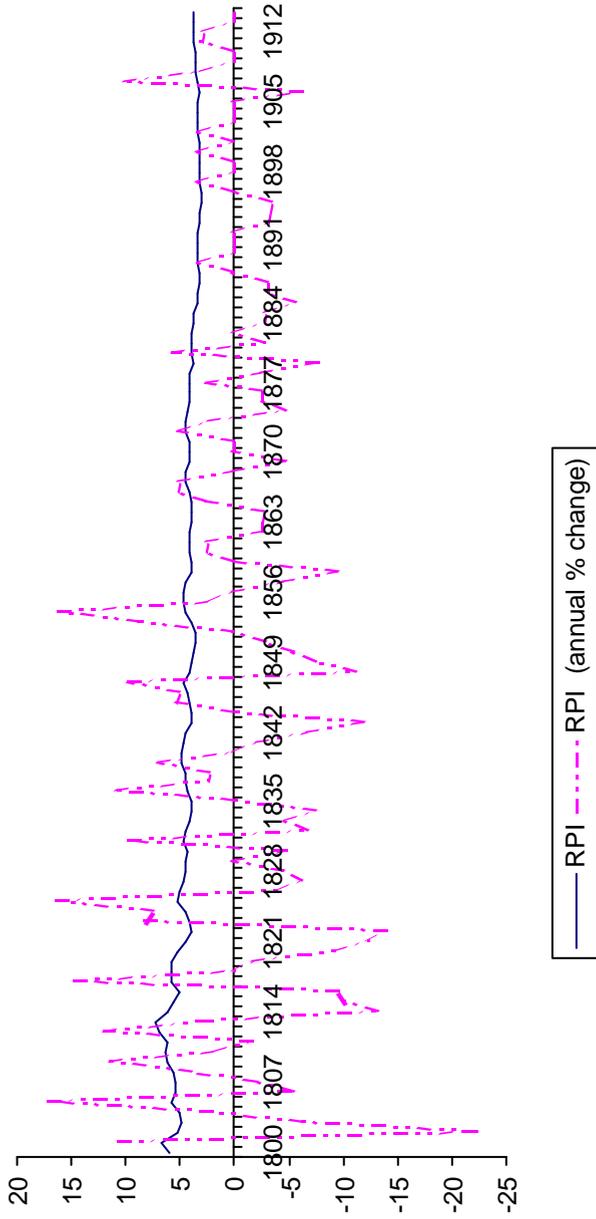


Chart 7

Bank Rate, inflation and £/\$ exchange rate, 1817-1914

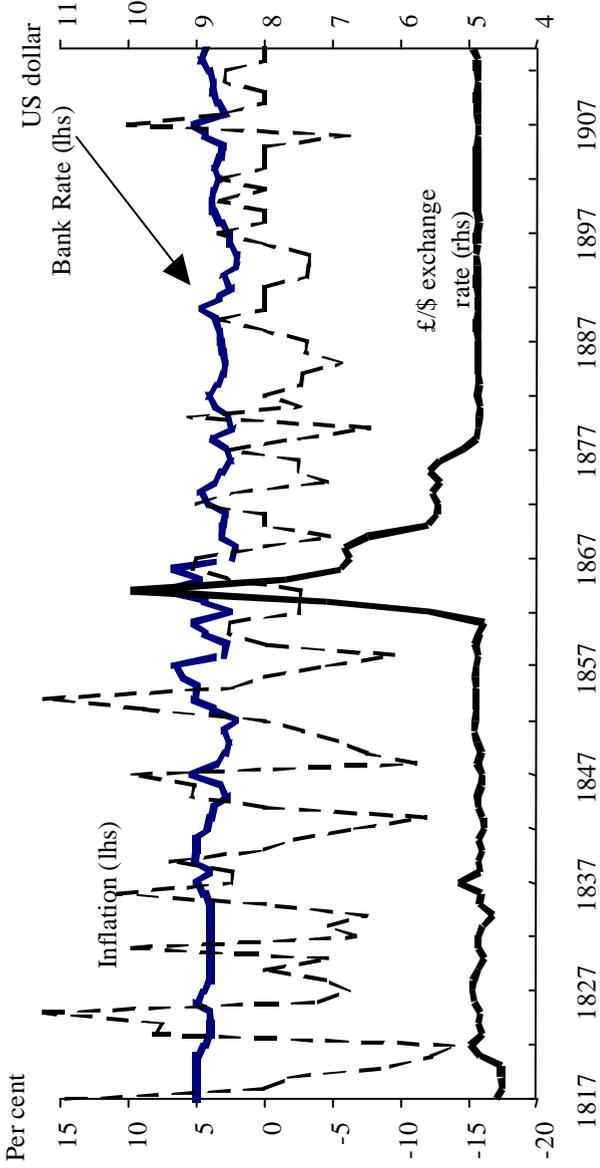
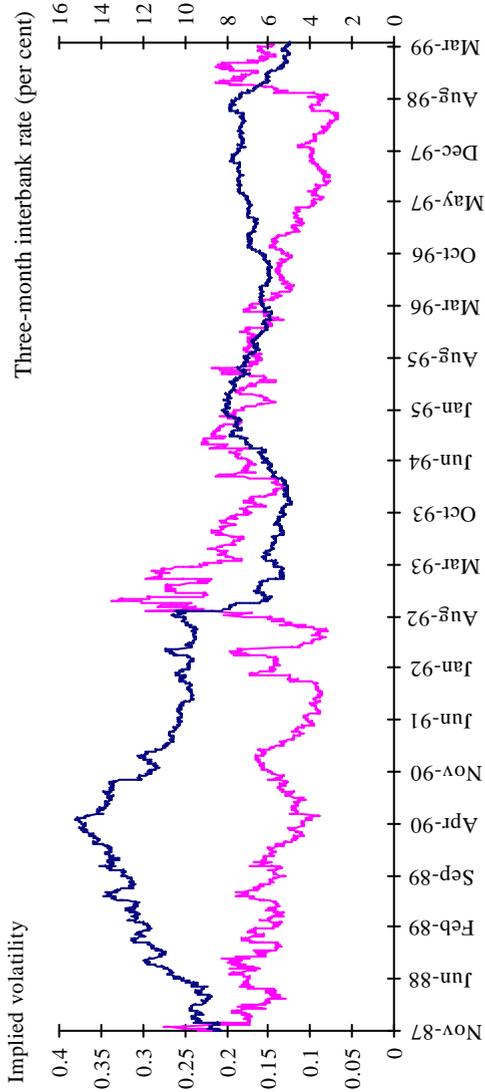


Chart 8

Short-sterling futures options (LIFFE) - constant horizon of six months



correlation between 1987-99: -0.26, 95% confidence intervals: -0.22 -0.30

correlation between 1993-99: -0.41, 95% confidence intervals: -0.35 -0.46

Chart 9
UK Base Rate 1800-1999

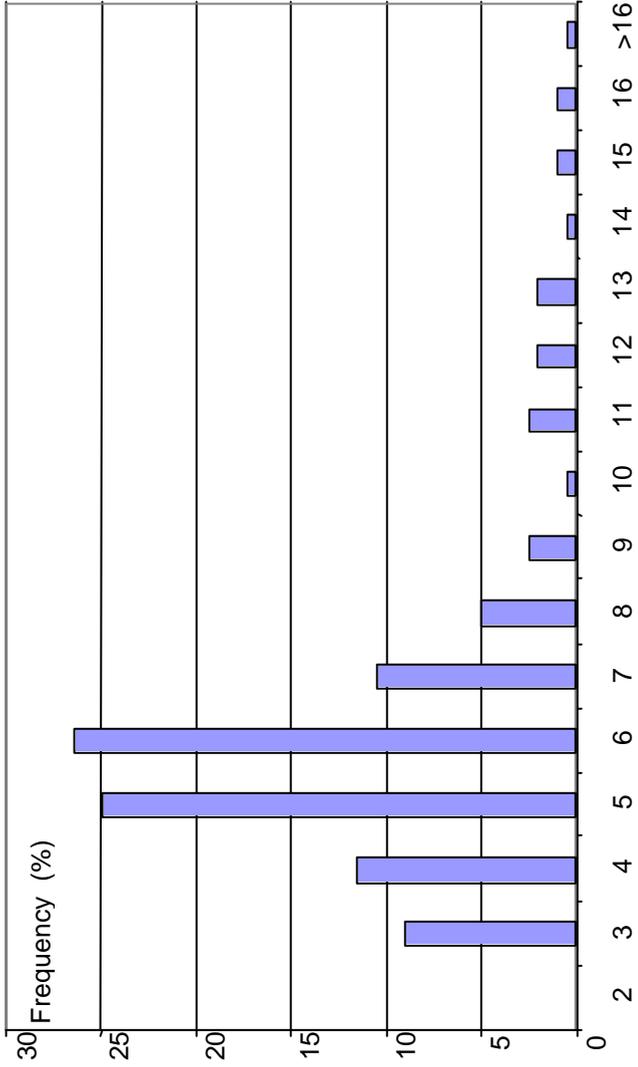


Chart 10

Histogram of standardised residuals: UK three-month interbank rate 1975-99

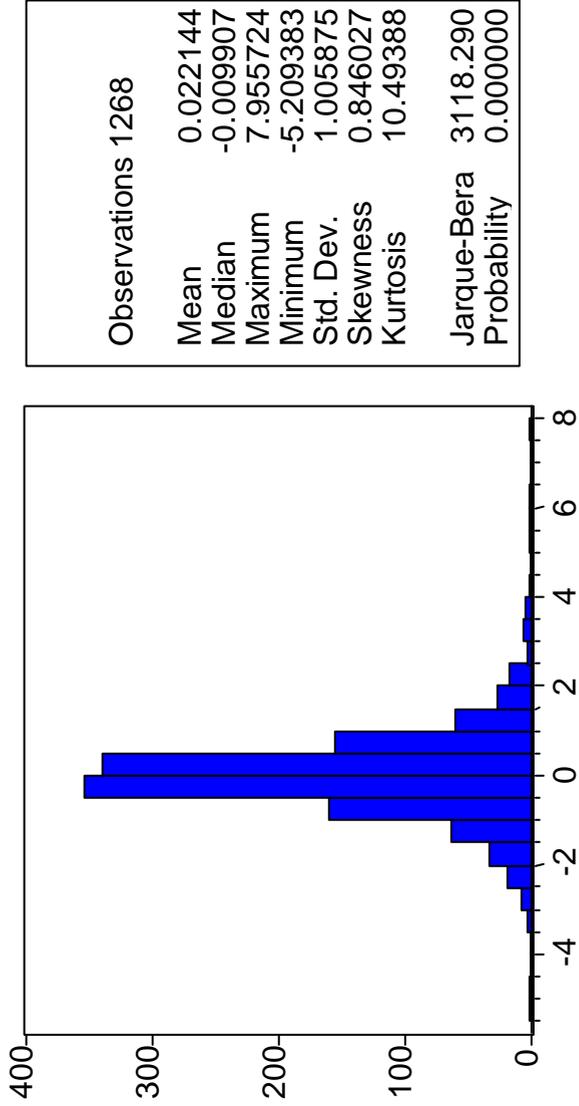


Chart 11

Histogram of standardised residuals: UK Base Rate 1800-1998

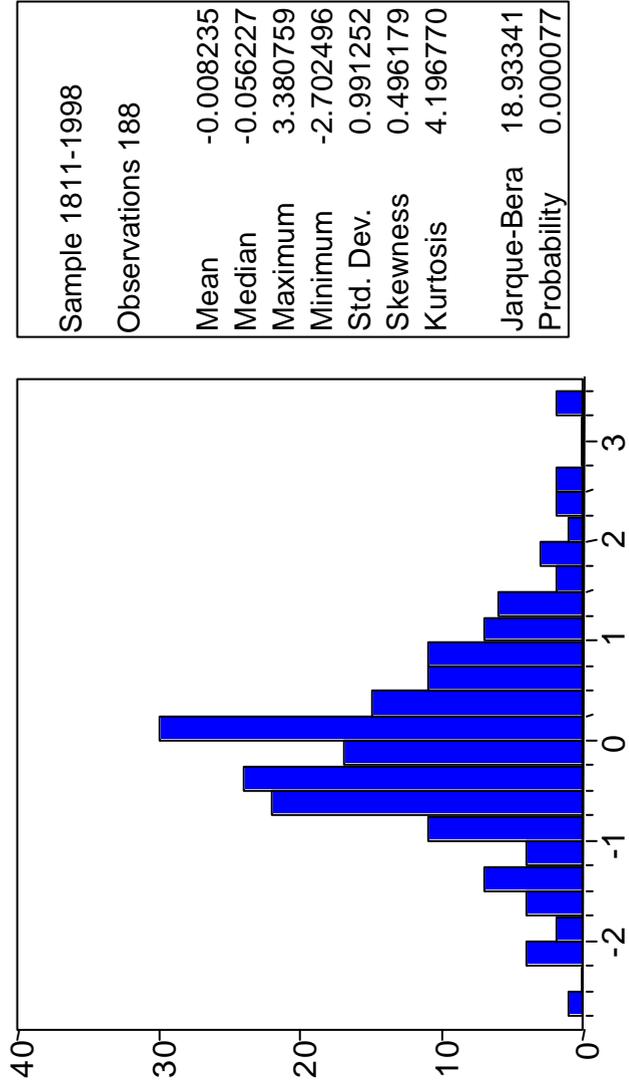
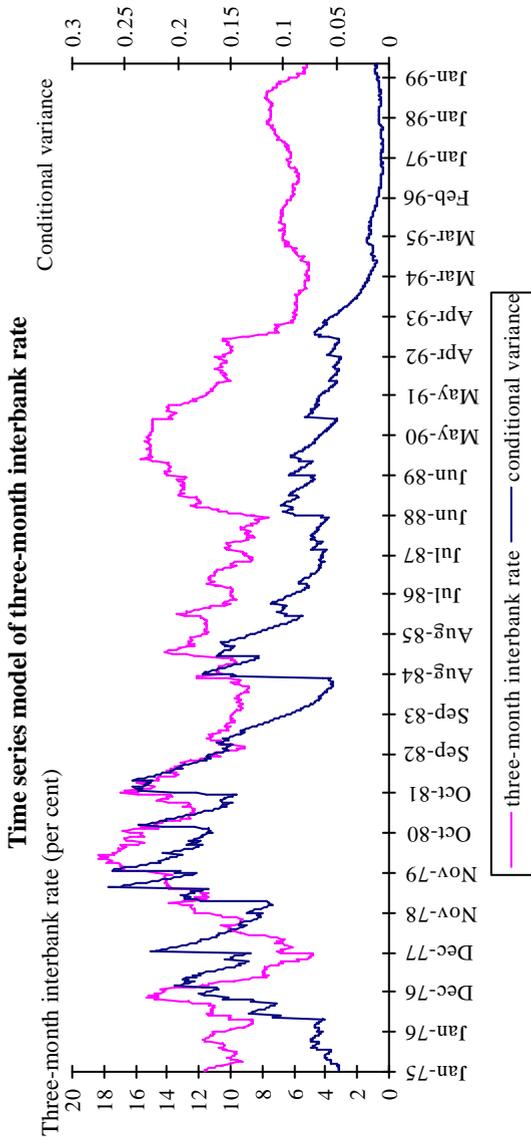


Chart 12

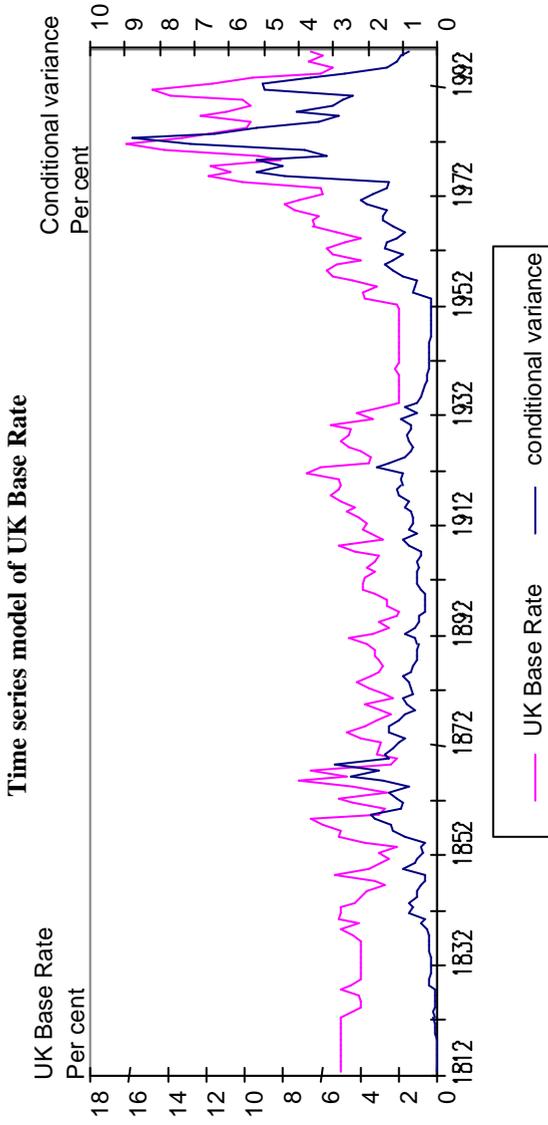


correlation between 1975-99: 0.66, 95% confidence intervals: 0.72 0.59

correlation between 1993-99: -0.26, 95% confidence intervals: -0.13 -0.38

steady-state three-month interbank rate: 7.55, 95% confidence intervals: 4.02 14.41

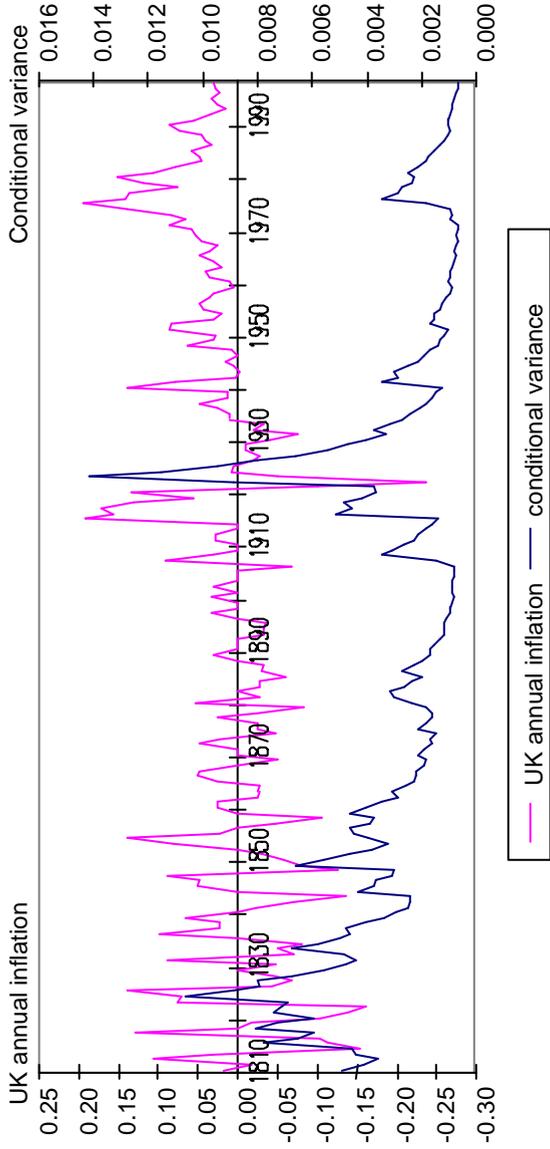
Chart 13



steady-state Base Rate: 4.88%, 95% confidence intervals: 1.02% 10%
correlation 0.81, 95% confidence intervals: 0.66 0.97

Chart 14

Time series model of UK annual inflation



steady-state inflation 2.7%, 95% confidence intervals: -10.7% 21.0%
correlation: -0.27, 95% confidence intervals: -0.43 -0.12

Chart 15

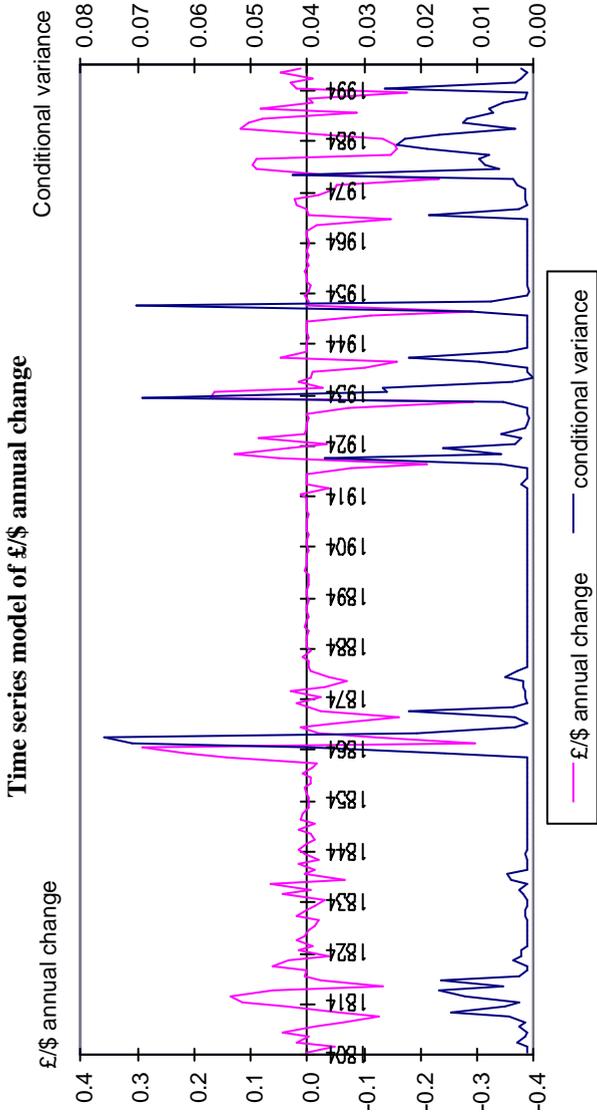
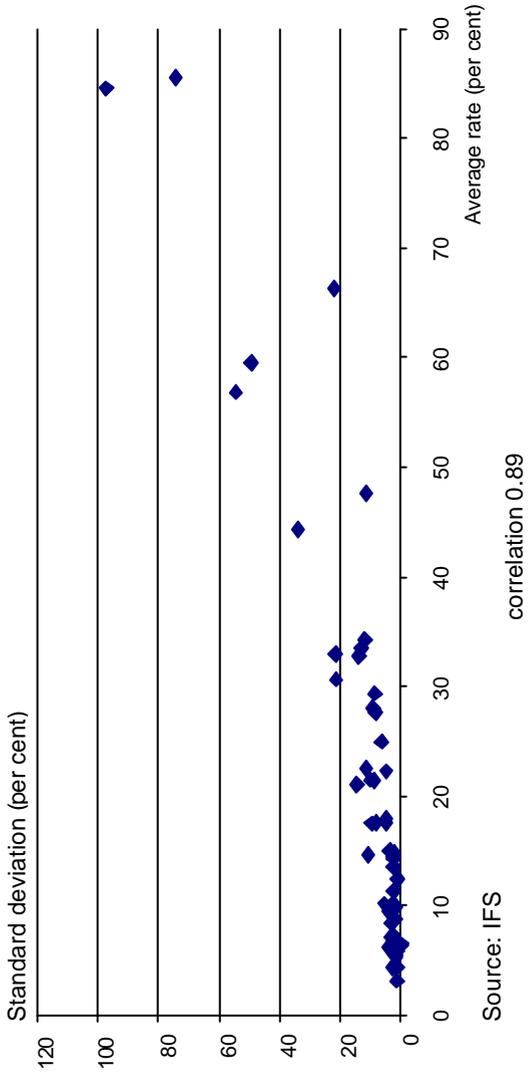


Chart 16

Scatter plot of the standard deviation and average level of treasury bill rates in 59 countries, January 1988 - January 1999



Appendix: Time-series investigation of the association between the level and volatility of the short nominal interest rate, the inflation rate and the exchange rate depreciation rate

(1) UK three-month interbank rate 1975-99

Let i denote the UK three-month interbank rate. The time series model estimated for Chart 1 was:

$$\begin{aligned}
 i_t - i_{t-1} &= a + bi_{t-1} + \mathbf{e}_t, \\
 \mathbf{e}_t &= \mathbf{r}\mathbf{e}_{t-1} + u_t + \mathbf{J}u_{t-1}, \\
 E(\mathbf{e}_t | \Psi_{t-1}) &= 0, E(\mathbf{e}_t^2 | \Psi_{t-1}) = \mathbf{s}_t^2, \\
 \mathbf{s}_t^2 &= \mathbf{w} + \mathbf{b}\mathbf{s}_{t-1}^2 + \mathbf{g}\mathbf{e}_{t-1}^2 + \mathbf{d}i_{t-1}
 \end{aligned} \tag{A1}$$

This time series model includes $AR(1)$ and $MA(1)$ terms in the conditional mean equation to account for the autocorrelation of standardised residuals. It includes $GARCH(1,1)$ terms in the conditional variance equation to account for the autocorrelation of squared standardised residuals, as well as a linear function of the lagged interest rate level to capture the potential dependence of the conditional variance on the lagged interest rate level.

Weekly data of the UK three-month interbank rate from January 1975 to April 1999 were used for the model estimation. The estimation results using maximum likelihood are shown in Table A.1.

Table A.1

	Coefficient	Std. error	Prob
<i>a</i>	0.047	0.030	0.119
<i>b</i>	-0.0062	0.0038	0.103
<i>r</i>	0.86	0.06	0.000
<i>J</i>	-0.78	0.08	0.000
<i>w</i>	-0.00042	0.00036	0.242
<i>b</i>	0.975	0.007	0.000
<i>g</i>	0.022	0.008	0.005
<i>d</i>	0.000069	0.000059	0.243

R^2 (adjusted) = 0.016

Standard errors are estimated using quasi maximum likelihood. The estimates of the parameters are consistent even if the conditional normality assumption is violated. They can, however, be inefficient.

The coefficient *d*, that determines the dependence of the conditional variance on the lagged interest rate level is insignificant. Alternative models of the conditional variance that include higher-order powers or higher-order lags of the lagged interest rate level were also estimated.⁽³⁷⁾ The coefficients of the higher-order powers or higher-order lags of the interest rate level were also insignificant. Therefore, the dependence of the conditional variance on past interest rate levels is weak.

However, the ex-post contemporaneous relationship between the interest rate level i_t and the conditional variance s_t^2 is strong. In fact the two variables have a correlation of 0.66. This is mainly because of the large information

(37) This includes a version replacing the conditional variance equation in (A1) by $s_t^2 = w + bs_{t-1}^2 + gs_{t-1}^2 + di_{t-1}^2$, a specification suggested by Chan, Karolyi, Longstaff and Sanders (1992).

shocks in 1970s and 1980s when short interest rates were high. After the introduction of inflation targeting in the United Kingdom (in late 1992), the relationship between the interest rate level i_t and the conditional variance \mathbf{s}_t^2 is weaker. In fact they are slightly negatively correlated, with a correlation coefficient of -0.26.

The steady-state forecast of the level of the three-month interbank rate is 7.546. The 95% confidence intervals are 4.02 and 14.41. The confidence intervals were constructed by assuming that the standardised steady-state forecast follows the in-sample distribution of the standardised residuals which, of course, has finite support. In particular the assumed skewness and kurtosis were 0.846 and 10.49 respectively.

Using $\ln(1 + i_t)$ rather than i_t as the specification of the interest rate variable in the regressions did not result in significantly different results.

(2) *UK base rate 1800-1998*

Let i denote the UK base rate. The time series model estimated was an EGARCH model of the form:

$$\begin{aligned}
 i_t - i_{t-1} &= a + bi_{t-1} + \mathbf{e}_t, \\
 \mathbf{e}_t &= \mathbf{r}\mathbf{e}_{t-10} + u_t + \mathbf{J}u_{t-2}, \\
 E(\mathbf{e}_t | \Psi_{t-1}) &= 0, E(\mathbf{e}_t^2 | \Psi_{t-1}) = \mathbf{s}_t^2, \\
 \log(\mathbf{s}_t^2) &= \mathbf{w} + \mathbf{b} \log(\mathbf{s}_{t-1}^2) + \mathbf{a} \left| \frac{\mathbf{e}_{t-1}}{\mathbf{s}_{t-1}} \right| + \mathbf{g} \frac{\mathbf{e}_{t-1}}{\mathbf{s}_{t-1}}
 \end{aligned} \tag{A2}$$

The coefficients \mathbf{a} and \mathbf{g} capture the potentially asymmetric impact of last periods shocks on conditional variance.

The time series model also includes $AR(10)$ and $MA(2)$ terms in the conditional mean equation to account for the autocorrelation of standardised residuals.

Annual data for the UK base rate from 1800 to 1998 were used for estimation. The estimation results using maximum likelihood are shown in Table A.2.

Table A.2

	Coefficient	Std. error	Prob
<i>a</i>	0.34	0.11	0.002
<i>b</i>	-0.071	0.023	0.002
<i>r</i>	0.12	0.07	0.099
<i>J</i>	-0.28	0.08	0.000
<i>w</i>	-0.064	0.086	0.455
<i>g</i>	0.29	0.07	0.000
<i>b</i>	0.91	0.02	0.000
<i>a</i>	0.078	0.102	0.446

$$R^2 \text{ (adjusted)} = 0.120$$

Standard errors are estimated using quasi maximum likelihood. The estimates of the parameters are asymptotically consistent even if the conditional normality assumption is violated. They can be inefficient, however, in this case.

The EGARCH model was chosen over alternative GARCH models, because its long-run unconditional variance was non-explosive.⁽³⁸⁾ The coefficient *g* is significant,⁽³⁹⁾ but the effect is the opposite of the ‘leverage’⁽⁴⁰⁾ effect, ie a

(38) An asymmetric component ARCH model had also a non-explosive unconditional variance, but the convergence was much slower than EGARCH model.

(39) At the 95% confidence level.

(40) The ‘leverage’ effect is the negative correlation between current returns and future volatility, found mainly in stock returns data.

negative shock has a negative impact on the conditional variance. The steady-state forecast of the base rate is 4.88. The 95% confidence intervals are 1.02 and 10.1. The confidence intervals were constructed by assuming that the standardised steady-state forecast follows the in-sample distribution of the standardised residuals. In particular the assumed skewness and kurtosis were 0.49 and 4.197 respectively.

(3) *UK annual inflation (RPI annual percentage changes) 1800-98*

Let p denote the RPI and $\mathbf{p}_t = \frac{p_t - p_{t-1}}{p_{t-1}}$ its annual proportional rate of

change. The time series model estimated was:

$$\begin{aligned}
 \mathbf{p}_t - \mathbf{p}_{t-1} &= a + b\mathbf{p}_{t-1} + \mathbf{e}_t, \\
 \mathbf{e}_t &= \mathbf{r}_1\mathbf{e}_{t-5} + \mathbf{r}_2\mathbf{e}_{t-8}, \\
 E(\mathbf{e}_t | \Psi_{t-1}) &= 0, E(\mathbf{e}_t^2 | \Psi_{t-1}) = \mathbf{s}_t^2, \\
 \mathbf{s}_t^2 &= \mathbf{w} + \mathbf{b}\mathbf{s}_{t-1}^2 + \mathbf{g}\mathbf{e}_{t-1}^2
 \end{aligned} \tag{A3}$$

This time series model includes *AR(5)* and *AR(8)* terms in the conditional mean equation to account for the autocorrelation of standardised residuals. It includes *GARCH(1,1)* terms in the conditional variance equation to account for the autocorrelation of squared standardised residuals.

Annual data of the UK RPI annual proportional changes from 1800 to 1998 were used for the model estimation. The estimation results using maximum likelihood were:

Table A.3

	Coefficient	Std. error	Prob
<i>a</i>	0.013	0.006	0.031
<i>b</i>	-0.46	0.085	0.000
<i>r</i> ₁	0.15	0.08	0.053
<i>r</i> ₂	0.24	0.08	0.001
<i>w</i>	0.00009	0.000103	0.380
<i>b</i>	0.81	0.12	0.000
<i>g</i>	0.16	0.104	0.114

$$R^2 \text{ (adjusted)} = 0.294$$

Standard errors are estimated using quasi maximum likelihood. The estimates of the parameters are asymptotically consistent even if the conditional normality assumption is violated. They can be inefficient, however, in this case.

Alternative models of the conditional variance that include inflation rate dependence were also estimated. The coefficient of the inflation rate dependence was insignificant.

The steady-state forecast of the inflation rate is 2.7%. The 95% confidence intervals are -10.7% and 21%. The confidence intervals were constructed by assuming that the standardised steady-state forecast follows the in-sample distribution of the standardised residuals. In particular, the assumed skewness and kurtosis were 0.434 and 5.286 respectively.

(4) £/\$ annual changes 1800-1998

Let *s* denote the annual spot exchange rate and $e_t = \frac{s_t - s_{t-1}}{s_{t-1}}$ its

proportional rate of change. The time series model estimated was:

$$\begin{aligned}
e_t - e_{t-1} &= a + br_{t-1} + \mathbf{e}_t, \\
\mathbf{e}_t &= \mathbf{r}\mathbf{e}_{t-2} + u_t + \mathbf{J}u_{t-3}, \\
E(\mathbf{e}_t | \Psi_{t-1}) &= 0, E(\mathbf{e}_t^2 | \Psi_{t-1}) = \mathbf{s}_t^2, \\
\mathbf{s}_t^2 &= \mathbf{w} + \mathbf{b}\mathbf{s}_{t-1}^2 + \mathbf{g}e_{t-1}^2 + \mathbf{d}e_{t-1}^2
\end{aligned} \tag{A4}$$

This time series model includes *AR*(2) and *MA*(3) terms in the conditional mean equation to account for the autocorrelation of standardised residuals. It includes *GARCH*(1,1) terms in the conditional variance equation to account for the autocorrelation of squared standardised residuals, as well as a linear term in the square of the growth rate of the exchange rate to capture the potential dependence of the conditional variance on the proportional rate of change of the exchange rate. Annual data on annual percentage changes of the £/\$ exchange rate from 1800 to 1998 were used for the model estimation. The estimation results using maximum likelihood were:

Table A.4

	Coefficient	Std. error	Prob
<i>a</i>	-0.0059	0.0016	0.000
<i>b</i>	-0.75	0.09	0.000
<i>r</i>	-0.18	0.05	0.000
<i>J</i>	-0.25	0.08	0.001
<i>w</i>	0.00101	0.00027	0.000
<i>b</i>	0.10	0.06	0.113
<i>g</i>	-0.13	0.06	0.034
<i>d</i>	0.89	0.32	0.006
R^2 (adjusted) = 0.422			

The coefficient *d*, which measures the dependence of the conditional variance on the squared proportional exchange rate change, is highly significant. The steady-state forecast of the £/\$ annual percentage change is -0.8%. The 95% confidence intervals are -10.6% and 4.2%. The confidence intervals were constructed by assuming that the standardised steady-state

forecast follows the in-sample distribution of the standardised residuals. In particular the assumed skewness and kurtosis were 0.864 and 7.57 respectively.

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